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Report of

Joint ICES/ICNAF Working Group on Selectivity Analysis

Comparative Selectivity of Bottom Trawls made of Different Materials

(Moscow, March 1969, and Charlottenlund, September 1969)

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INTRODUCTION

Following the report of its Subcommittee on Fishing Gear and Selectivity, the Standing Committee on Research and Statistics (R&S) at the 18th Annual Meeting of ICNAF held in London from 4 to 8 June 1968 recommended:

- that an ICNAF Working Group on Selectivity Analysis be formed to undertake a review of variability in selection data, including the scientific basis of mesh size differentials for different twine materials;
- that the Chairman of R&S appoint a convener for this Working Group; **ii**)
- iii) that experts be appointed to the Working Group by interested member countries;
- that ICES be invited to participate, in order to make a joint study of these problems; iv)
- that the Working Group meet at a mutually convenient time so that its report could be available for the next meeting of both ICNAF and NEAFC.

(ICNAF Redbook 1968, Pt.I, p.12)

At its 56th Annual Meeting held in Copenhagen from 30 September to 9 October 1968, ICES accepted ICNAF's invitation to take part in the activities of the Group (C.Res. 1968/3:3).

A preparatory meeting was held in Copenhagen on 4 October 1968 attended by representatives of those countries which expressed their wish to participate in the activities of the Group, at which time, place and program of work were agreed upon.

The Joint ICES/ICNAF Working Group on Selectivity Analysis held its meetings at VNIRO, Moscow, from 26 to 29 March 1969 and at ICES, Charlottenlund, from 2 to 6 September 1969.

The Working Group was to prepare working papers covering the following subjects:

- Comparison of the properties of net materials of trawls in the North Atlantic and investigation of the effect of trawl construction on selectivity;
- 2) Analysis of the variability of marine experiments on selectivity and the validity of selectivity data;
- Compilation of the selectivity data for cod, haddock and redfish, including a tabulation of equivalents for different net materials.

Prof. von Brandt was asked to prepare the first working paper, Mr Pope the second, and Dr Treschev the third. These three papers appear as Parts A, B, and C of this report. A bibliography was prepared by Dr Bohl and Mr Holden which appears as Part E.

The meeting participants were as follows:

Both meetings

Dr A.I.Treschev, USSR (Chairman)

Mr M.J.Holden, UK (Secretary)

Mr J.A.Pope, UK Dr H. Bohl, FRG Prof. A. von Brandt, FRG

First meeting only

Dr J. Reuter, Netherlands

Mr S. Prüffer, Poland

Dr W. Strzyzewski, Poland

Dr G.N.Stepanov, USSR

Mr B.I.Danilov, USSR

Mr A.A.Volkov, USSR

Second meeting only

Mr V.S.Belov, USSR

Mr L.M.Zheltov, USSR

Mr J. Møller Christensen, Secretary of the Liaison Committee

PART A

COMPARATIVE EVALUATION OF THE PROPERTIES OF NET MATERIALS USED FOR MAKING TRAWLS

by
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According to different investigations, the selectivity of codends in trawls may be influenced by different facts:

- a) fish (species, quantity, behaviour);
- b) gear (material, rigging, hanging, mesh size, shape of the mesh);
- c) operation of the gear (speed, towing force, towing time).

Most investigations underline that the net material has a decisive influence on the selectivity of a trawl. However, there are also examples, e.g. size of the catch or shape of the mesh, that may have an influence.

The selection factor, sf., that means the relation between fish length and mesh size, is considered to be more or less constant for the different net materials used for codends in bottom trawling. According to the results of many tests for the same material (e.g. polypropylene), the construction (e.g. twisted or plaited) and the type of material (e.g. fibre or film) have no influence.

It has been found that different net materials used for codends in the ICNAF and ICES areas have similar or the same selection factor. There are at least two groups:

- 1) manila, sisal, polyethylene (PE), and polypropylene (PP);
- 2) cotton, hemp, polyamide (PA), and polyester (PES).

Nevertheless, sometimes doubt arose about the classification of some materials. There may be some overlapping according to the complexity of the problem.

The different selectivity of different net materials used for codends in trawling has been explained by other authors with the different properties of the net materials. Extension, flexibility, and roughness have been mentioned as the main reasons for different selectivity.

The difficulty is that the properties of a netting twine are changeable in some degree with the type of manufacture. A hard twisted twine can have not only less strength but also less flexibility compared with a soft twisted one. The tension of netting twines increases with their numbers or diameters, etc. This means that there must be more than one property responsible for the selectivity, even when the net material is taken into consideration.

In the case where different net materials have almost the same selectivity as mentioned before, and their properties are responsible for the selectivity, they must have similar properties. To find them could give an explanation for the theory of the selection process in codends.

Only those net materials which can be exchanged with each other are comparable. According to the differences of properties, it seems unwise to compare net materials used for inshore fishing with small boats with those used for deep sea fishing with big trawlers. Therefore, member countries have been asked for materials used for bottom trawls in the ICES or ICNAF areas only. It may be useful to limit the net materials to R 2,500 tex and 400 m/kg.

The following countries have been kind enough to send samples of net materials: Belgium; Canada; Denmark; Finland; France; Germany, Fed. Rep.; Iceland; Ireland; Italy; Netherlands; Poland; Portugal; Romania; Spain; Sweden; UK; USA; USSR. Samples are missing from Norway only.

The data for the properties of the netting twines are given in Tables A-1 to A-7. These tables show the local designations of the materials and the real ones in R.tex and m/kg (yds/lb.). Moreover, the following properties are included: diameter, knot breaking load, elongation at a load corresponding to the half knot breaking load in percent and the flexibility. The testing methods follow the rules under discussion in the International Organization for Standardization (ISO), Technical Committee 38/Textiles, Subcommittee 9: Textile Products for Fishing Nets. For the test of the flexibility, a method not yet accepted has been used in spite of some objections made in a previous ICES paper 10 years ago.

A preliminary comparison between the properties of the net materials used for codends and the two groups of similar selectivity mentioned above show some relation with the flexibility (Fig. A-1). The net materials made of PA, PES, and also unexpectedly PP have almost the same range of flexibility in contrast to net materials made of manila and PE, though there are some extreme values outside the main range, shown by dotted lines.

Moreover, a diagram with the relation of tension and flexibility shows on one side PA and PES and on the other manila and PE with some overlapping by the data for PP in both areas (Fig. A-2).

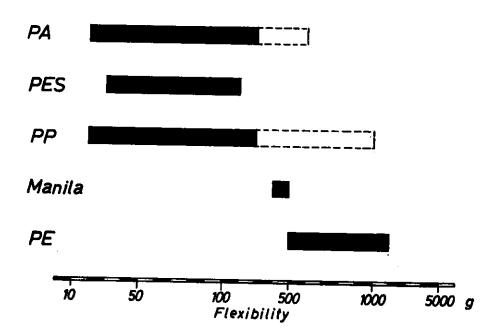


Fig. 1. Comparison of flexibility of net materials used for codends.

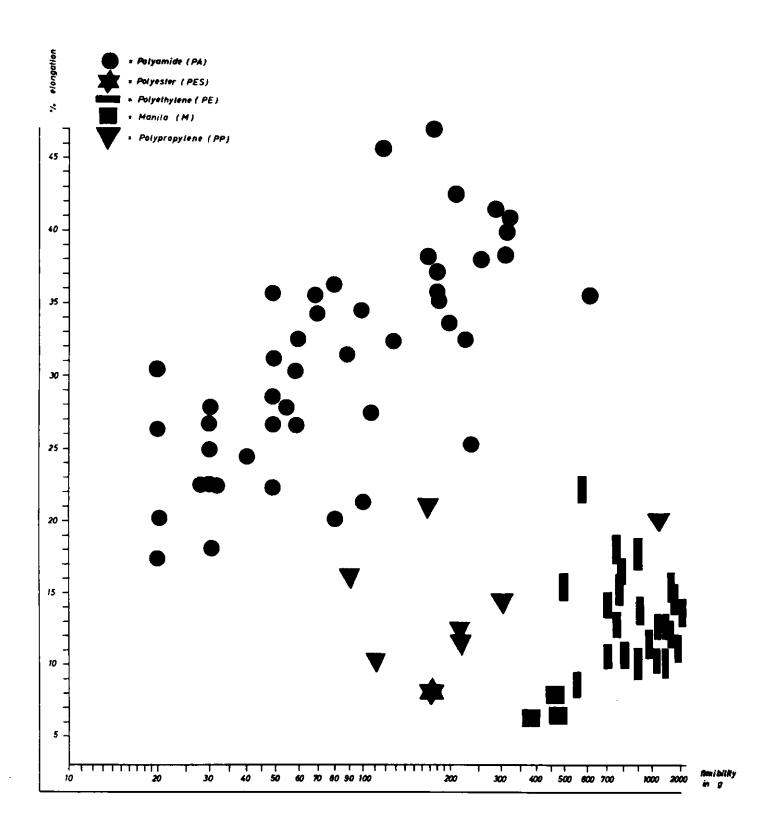


Fig. 2. Elongation/flexibility (only values > R.2500 tex or <400 m/kg).

Remarks orange green orange " Flexi-bility (g) 72 10 10 10 20 20 50 50 50 50 130 180 170 250 220 270 270 370 580 700 800 ponding to the half knot a load corres-Elongation at breaking load 20.2 28.8 29.7 29.7 31.8 30.3 31.2 31.2 26.9 32.4 32.4 11.0 12.0 13.2 15.4 12.3 10.5 10.6 Knot breaking load, wet 30.4 29.9 37.6 41.4 45.9 52.9 76.9 93.4 116.6 34.9 58.9 65.4 105.6 106.3 104.0 130.4 construct- Diameter (E) 0.75 0.91 1.00 1.11 1.26 1.44 1.68 1.84 1.84 1.85 2.22 1.10 1.65 1.68 1.65 1.78 2.06 2.27 2.27 twisted Twine ion yds/1b. Runnage 2,100 1,700 1,300 1,050 850 700 500 450 350 360 200 1,200 700 650 550 500 450 350 300 11/kg R.tex (g/1,000 m) 500 600 750 950 1,150 1,450 2,000 2,300 2,850 4,800 850 1,400 1,550 1,800 2,000 2,150 2,850 3,250 Polyethylene (PE) Belgium: Polyamide (PA) number Trade

Table A-1.

Table A-2.

Trade	R.tex (g/1,000 m)	Runnage m/kg	yds/1b.	Twine construct- ion	Diameter (mm)	Knot breaking load,wet (kp)	Elongation at a load corres- ponding to the half knot breaking load (2)	Flexi- bility (g)	Remarks
Canada: Polyethylene	ene (PE)								
4.5 mm	6,100	150		plaited	(4.5)	242	10.3	1,550	orange
	6,400	150		: =	(4.5)	239	12.5	1,590	Į,
5.5	8,950	007		=	(5.5)	777	13.1	2,140	orange
5.5 mm co_oo	5,100	700 200		Ξ	(3.0)	208	10.6	1,180) =
CP-160	3,250	300		=		164	15.5	780	=
Terylene	(PES)								
T-50	7,200	150	75 (70)	=		288	8.5	170	
Nylon (PA)	ন								
N-50	5,800	150	75 (75)	=		288	31.6	90	
Ulstron	(PP)								
	3,850	250		: :		237 402	10.4	110 220	
UP- 76	7,450	150		Ξ		393	11.3	220	
	6,950	150		I :		389	14.9		
3/16 in.	14,500	50		=		676	20.3	1,140	golden, with core
Denmark: Polyamide	le (PA)								
Td 210/24	009	1,600		twisted "	0.77	39.3	16.0 16.2	7 V V	
	950	1,050		E	96.0	61.8	17.9	S	
	5,350	200		plaited	(3.0)	242	26.8	60	
4 mm	7,050	150		:	(4.0)	920	5. /7	077	

Values shown in parentheses are nominal and/or given by the factories.

		Remarks					treated black																			-							
	Flexi-	bility (g)		-2	\. \.	2	ır	20	40		0,	3 8	20	100	290	7 7 7 0 0	T,490		< 2	7.2	< 2	20	09	20	30	30	30	20	80	740		700	890
Elongation at a load corres-		breaking load (%)		17.3	17.8	23.2	16.6		22.1		20.7	25.0	28.4	21.3	41.7	72.4	44.0			20-25	67 _07		29		!	20-25			20	22		14.1	17.5
4 1 2	knot breaking	load,wet (kp)		22	31	67	22	33	99		13/	194	291	285	494	6/0	0.68		32	38	41	06	80-85	100	140-152	160	2/0	340-390	780	111		228	352
		Diameter (mm)									(7 6)			(4.5)	(5.5)		٥						7			_			(5.0)	(7.0)		(4.0)	(2.0)
				0.52	0.67	1.05	0.69	0.88	1.40				3,53				76.6		0.80	0.84	0.90	1.50	1.6-1.7	1.65	2.00	$\frac{2.10}{2.10}$	3.00	3.50					
	Twine	construct-		twisted	=	=	Ξ	Ξ	=		בים + המן בי	1 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1	twisted	plaited	= =	P()	rwisted		=	Ξ	Ξ	Ξ	z ;	= :	: :	: =	: =	: 1	plaited	:		= :	=
		Runnage m/kg yds/1b.		3,300(3,300)	2,200(2,210)	1,000(1,100)	2.950	1,850	800		350	250	150	150	50	000	2		1,800	1,550	1,400	650	550	500	350	300	150	00T	100	90		200	150
		R.tex (g/1,000 m)	(PA)	300(320)	450(480)	1,000(960)	350(320)	550(480)	1,200(960)	de (PA)	2.850 (2.857)			(%) (100 (%)	13,550(11,363)	24, 700 (21, 276)	() (1 (1 (1 (1 (1 (1 (1 (1 (1 (1 (1 (1 (1	ide (PA)	550	650	200	1,600	1,800	2,000	3,050	3,400	6,230	. 300 .	11,000	18,000		4,650	7,550
		Trade	Finland: Polyamide	Td 210/12	Td 210/18	Td 210/36	Td 210/12		Td 210/36	France: Polyamide	F11. 2.4 6			. 4.5	5.5	F11: 0 W	Þ	Germany: Polyamide	Td 210/21									3/	5 mm 6	97//	<u>Polyethylene</u>	4 mm \$	III II

Table A-3

Values shown in parentheses are nominal and/or given by the factories.

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Tab
H

						Knot	Elongation at a load corres- ponding to the	 	
	ſ	1		Twine		breaking	half knot	Flexi-	
Trade number	R.tex (g/1,000 m)	Runnage m/kg	yds/1b.	construct- ion	Diameter (mm)	load,wet (kp)	breaking load (%)	bility (9)	Remarks (p=producer)
Holland: Polyamide	ide (PA)							Na Carlo	
				twisted	2.64	185.2	37.2	180	Λ
Td 210/132	3,800 (3,800)	250 (262.2)			2.30	164.4	30.2	20	- H
				=	2.13	163.6	26.4	20	. I Q
				=	2.20	155.5	34.5	70	II d
		250 (266.4)		=	2.49	151.6	38.2	170	III
				=	2.36	157.3	34.8	100	
Td 210/108				=	2.11	142.9	35.5	70	
				=	2.11	123.8	30.5	09	
				=	2.25	148.5	36.6	80	Δ Q
Td 210/96				=	1.83	130.9	23.8	10	
				=	2.00	141.6	35.8	20	
		300 (317.7)		=	2.26	127.7	42.6	210	III
	3,050 (3,000)			E	2.12	131.9	32.7	09	
Iceland: Polyethylene (PE)	hylene (PE)								
Ø mm 7	5,950	150	-	plaited	(4.0)	260	12.6	750	T or good o
5 tm 6	7,400	150		=	(2.0)	323	13.7	910	1 L L L
7 mm 7	5,800	150		=	(0.4)	217	17.9	740	s o
_	5,200	200		=	(4.0)	229	12.0	1,170	. D
S mm 6	8,150	100		Ξ	(2.0)	302	10.9	1,740	<u>-</u>

Values shown in parentheses are nominal and/or given by the factories.

8 K			ge		98		98			. X-											E						ជ
Remarks		red	orange	red	orange	red	orange		,	black "											green						green
Flexi- bility (g)		890	066	790	580	240	200		15	30 30 30			55		δ	30	20	20	30 40 40		1,650		310		630		1,700
Elongation at a load corres- ponding to the half knot breaking load		10.0	11.2	16.4	22.2	8.4	15.1		18.7	15.9 15.5			28.8		23.7	29.1	26.7	24.8	26.8 24.5		11.4		41.3		35.4		13.9
Knot breaking load,wet (kp)		184	171	162	147	133	153		69.1	37.4 31.8			173		31.0	71.4	87.7	108.9	146 253		290		261		377		318
Diameter (mm)		7.93	2.55		2.53	2.27	2.17		1.84	1.00 0.94									2.13 (2.5) 2.96 (3.0)		(5.0)		(5.5)				
Twine construct- ion		twisted		plaited	twisted	=	=		: ;	= =			plaited		twisted	Ξ	= :	= ;	: :		plaited		twisted		plaited		:
age vds/1b.		125 (117)	125 (135)				175 (190)		_	670 (780) 750 (890)																	
Runnage m/kg		750	250	300	350	350	350		550	1,350	<u>.</u>		250		1,700	009	550	450	300 150		150		59		100		100
R.tex (2/1.000 m)	(PE)	(300 (% 5%0)			2,900 (2,970)			e (PA)	U	750 (636) 650 (557)) }	(PA)	3,900	(PA)	(009) 009		1,850 (1,690)		3,350 (3,560) 6,550 (7,000)	ylene (PE)	7,850	e (PA)	17,055	(<u>PA)</u>	11,050	ne (PE)	8,150
Trade	Ireland: Polyethylene							Polyamide				Italy: Polyamide		Poland: Polyamide	Td 210/7x3				Td 840/2x5x3 Td 840/4x5x3	Portugal: Polyethylene	5 mm 9	Romania: Polyamide	5.5 mm Ø	Spain: Polyamide		Polyethylene	

dvalues shown in parentheses are nominal and/or given by the factories.

slippage or knot reated all green Remarks for knot slippage treated treated black black green green 380 460 460 Flexibility 30 60 90 170 10 180 (8) 30 70 140 120 180 3288 20 ponding to the a load corresbreaking load Elongation at half knot 6.5 7.8 6.7 18.4 35.6 8.3 9.2 12.3 12.4 45.8 16.5 17.5 27.8 18.0 10.5 16.8 16.1 21.3 8 load, wet breaking 46 (kp) 102 243 42 52 79 237 70 116 136 183 80 110 142 196 116 152 163 3.39 (3.6) 4.28 (4.33) (3.0)Diameter 1.57 1.33 1.21 1.44 1.76 1.41 1.73 2.66 2.82 1.53 2.30 2.57 3.19 2.57 3.15 3.18 construct-Ewisted twisted plaited Twine ion 125 (112.5) 100 (90) 75 (75) (290) (200) (150) (100) (290) (200) (150) (100) yds/1b 300 200 125 100 300 200 150 Runnage m/kg 500 150 950 700 400 1,100 150 600 400 250 200 650 400 300 200 250 200 150 1,600 (1,710) 2,550 (2,480) 3,900 (3,307) 5,050 (4,960) Polypropylene (PP) 1,500 (1,710) 2,600 (2,480) 3,450 (3,307) 5,450 (4,960) 7,000 (7,100) 10,750(10,860) 4,200 (4,409) 5,550 (5,512) 6,300 (6,614) (g/1,000 m) Great Britain: Polyamide (PA) R.tex Manila (Ma) Polypropylene (PP) 1,450 2,400 1,950 5,900 900 1,050 Polyester (PES) Sweden: Polyamide (PA) USSR: Polyamide (PA) Values shown in number Trade No.4:20 No. 10.7 No. 10.7 No. 10 No.14 No.20 No.36 3

parentheses are nominal and/or given by the factories.

Table A-6

Elongation at a load corres- Knot ponding to the	Twine breaking half knot F	$y_{\rm us/LD}$, $z_{\rm on}$ (mm) (kp) (%) (g) Remarks		(49.86) plaited (.187) a (40.28) " (.2187) a (
Knot	breaking load,we	(kp)		227 338 416 386
				2.93 (3.0) a (.187) (.2187) (.2187) (.187) a (.187) (.187) a (.187) a (.2187)
	Twine construct-	TOD		
	40 /11	yus/10.		50 (49.86) 50 (40.28) 50 (38.01) 25 (34.01)
	Runnage	m/ NB		150 100 100 100 50
	R.tex (0/1 000 m)	(m coot +)4)	(PA)	5,750 10,350 (9,948) 12,600 (12,313) 13,300 (13,050) 14,550 (14,558)
	Trade		USA: Polyamide (PA)	3x840x16

Table A-7.

 $\frac{a}{}$ = App. diameter given by USA

PART B

THE STATISTICAL ANALYSIS OF SELECTIVITY DATA

by
J.A.Pope
Marine Laboratory, Aberdeen, Scotland

1. Introduction

The length-selection curve of a trawl codend gives the relationship between the probability of retention of a fish by that codend and its length. Expressed in general mathematical terms, the form of the relationship is $P = f(1|\underline{\theta})$, where P is the probability of retention, f is some function of the length of fish (1) and $\underline{\theta}$ is a vector of parameters defining the characteristics of the length-selection curve.

From both theoretical considerations and practical evidence, the function $f(1 \mid \theta)$ is taken to be monotonic increasing with $f_{max} = 1$ and $f_{min} = 0$. Largely because of these mathematical properties, it has been generally accepted that $f(1 \mid \theta)$ can be realistically and accurately described by some continuous probability distribution function and, in particular, by the symmetric distribution function of the Normal (Gaussian) frequency distribution. Jones (1957) has shown how, on the basis of some specific assumptions concerning variation in codend mesh size and fish body girths, the length-selection curve is exactly expressible as a Normal distribution function. However, this result is entirely dependent on the assumed normality of the frequency distributions of both mesh size in the codend and fish body girth. Other assumptions would lead to other expressions for the

Pope and Hall (1960) expressed $f(1|\underline{\theta})$ as a logistic function which is also symmetric and, in fact, differs little from the Normal distribution function. These writers did not attempt nor imply any theoretical justification for this form of curve. A full description of the fitting of this form to experimental data is given by Pope (1964).

Both the Normal distribution function and the logistic function are completely specified by two parameters θ_1 and θ_2 (say) in terms of which all points on the selection curve (in particular the 50% selection length) may be determined.

There does exist, however, an extensive body of experimental data indicating that a selection curve is not necessarily symmetric about the 50% selection length. A more suitable and flexible mathematical model for the selection curve is clearly desirable.

In practice, however, interest is generally confined to the 50% selection point rather than to the curve as a whole. Little attention appears to have been paid to the selection range (i.e. the length interval between the 25 and 75% selection points). The most important practical problem has, therefore, been that of estimating 50% selection points from experimental results.

2. Estimation of a Single Selection Curve

Before attempting any estimation, consideration ought to be given to whether the data to be used are sufficiently reliable to warrant any form of treatment at all. Such consideration, of course, begs the questions, "what is to be estimated?" and "what accuracy is required in the estimates?" Although a complete selection curve is usually desired, the main requirement, as already to methods of achieving this.

First we must distinguish between various sources of statistical error. We consider the results from a single covered codend haul as being a single realization of how the codend selects fish under the conditions pertaining at the time of towing. These conditions refer to the composition of the fish population encountered, the flow of water through the codend during towing, the

speed of the vessel, etc. Another haul with the same codend will not be made under precisely the same conditions and it must be regarded as a single realization of another situation. Whether the two sets of conditions result in closely similar or widely dissimilar average selection curves depends simply on how similar the different conditions are on these two occasions. If $p_j(1_k)$, k =1, 2, 3, are the observed proportions retained at lengths 1_1 , 1_2 , 1_3 , ... in a particular haul, say haul j, and $f_1(1_k|\underline{\theta})$ the expected proportions derived on the basis of some theoretical law (e.g. that the probabilities of retention follow a Gaussian distribution function law), then we may write

$$p_{j}(1_{k}) = f_{j}(1_{k}|\underline{\theta}) + \epsilon_{jk}$$
 (2.1)

The fitting of $f_j(1_k|\underline{\theta})$ to the data from this single haul consists of finding estimates of the parameters $\underline{\theta}$. Such estimates $(\underline{\hat{\theta}}, say)$ will be functions of the statistical errors (ε_{jk}) and the best estimates are those which involve these errors in such a way as to make their generalized variance (given by $\det \{ E(\hat{\theta} - \theta)'(\hat{\theta} - \theta) \}$) as small as possible. Maximum likelihood estimators have this property in large samples, but, in general, it is not difficult to find, by simpler methods, estimates of the parameters, or functions of them, which have sufficiently small variance.

Whilst the Gaussian distribution function, $f_{\underline{G}}(1|\underline{\theta})$ say, or the logistic, $f_{\underline{L}}(1|\underline{\theta})$ say, are strictly at best only assumed forms for the function $f(1|\underline{\theta})$, these curves are, at least on empirical evidence, sufficiently realistic for them to be regarded as equally plausible "true" models in a great many situations. The fitting of either curve may thus be regarded as being the most advanced form of treatment of the data. The parameters of these curves may be estimated by maximum likelihood using suitable routines and their variances, or more often the variances of functions of them (e.g. their ratio), evaluated. For example, the logistic curve is given by

$$1/f_{L}(1|\underline{\theta}) = \left\{1 + \exp(\theta_{1} - \theta_{2} - \theta_{2})\right\}$$
(2.2)

and the ratio θ_1/θ_2 is the 50% retention length.

Any method of fitting other than maximum likelihood will not be more efficient, in the sense that it will lead theoretically to variances of the estimates as large as or larger than those of the maximum likelihood estimates. Also, any form of curve other than $f_L(1|\underline{\theta})$ or $f_G(1|\underline{\theta})$ can be regarded as an approximation to the true model and, like all approximations, will theoretically be less accurate than the true model. Nonetheless, the amount of calculation involved in the maximum likelihood approach and the need for special tables makes it very desirable to seek alternative simpler ways of treating the data.

Simpler approaches are

1)

curve fitting by eye approximating to the middle part of the curve by a straight line and fitting by 11) weighted or unweighted least squares

using moving averages of the observed values of $f_j(1_k)$ to estimate the 50% point only.

The results of two sets of comparisons of different methods of fitting curves are given in Tables B-1, B-2, B-3 and B-4. The first two tables are taken from Pope (1966) and give the estimated 50% points and variances respectively of 15 actual selection curves provided by (a) the logistic curve fitted by maximum likelihood, (b) the fitting of an unweighted regression line to the observed proportions retained over approximately the 25-75% range, (c) the 3-point moving average method. Tables B-3 and B-4 give similar data for 9 actual curves provided by methods (a), (b) and (c) above, and (d) eye-fitted curves. These sets of data were taken from two series of replicate hauls made by the same research vessel during two cruises. Within each set the results refer to the same codend. The average values of the variances of the maximum likelihood estimates of θ_1/θ_2 , the 50% retention

lengths, are 0.1023 and 0.0968. These figures imply a percentage standard deviation of between 1.1 and 1.3% on average in the maximum likelihood estimates of a 50% retention length (and hence of a selection factor) from a single set of data (e.g. single haul). This is a very small component of

The unweighted linear regression estimates are satisfactorily close to the maximum likelihood estimates and, although all but one of the estimates in Table B-3 are below the maximum likelihood values, there is no reason to suspect bias in this method. The variances of the regression estimates differ appreciably from those of the maximum likelihood estimates in some, but certainly

not all, cases. The average value of the regression variances in Tables B-2 and B-4 are 0.1758 and 0.1202 respectively, corresponding, on average, to percentage standard errors of 1.7 and 1.2% respectively, that is, slightly larger than the maximum likelihood estimates.

The estimates provided by the use of 3-point moving averages are, like the regression estimates, close to the maximum likelihood values with no evidence of bias. The variances are similar in some cases to those for the above methods but are generally higher, sometimes appreciably so.

The estimates from eye-fitted curves, given in Table B-3, were obtained by asking seven people to draw curves by eye, independently of one another, to each of the nine sets of points. None of the seven people had previously drawn a selection curve so that it was necessary to give them some guidance beforehand. This consisted of outlining the principles of curve fitting to any sort of data, giving a brief account of codend mesh selection and stating simply the properties of the logistic and Normal distribution function curves. Each person was asked to fit a curve to all the points over the whole range, not merely in the neighbourhood of the 50% point. Thus the deviations of each observed proportion from each eye-fitted curve could be measured and compared. The sums of squares of these deviations are given in Table B-5.

It will be seen immediately from the last column of Table B-3 which gives the averages of the seven eye estimates for each haul, that eye estimation can give 50% points very close to maximum likelihood estimates. What is interesting and of significance in this context is that differences between the seven estimates are small in each of the nine cases. This is reflected in the variances calculated from the seven estimates and shown in the last column of Table B-4. These variances are not estimating precisely the same thing as the variances in the first three columns of Table B-4. They do indicate the degree of reproducibility of eye estimates, however, which is the important point here.

We may conclude from the results of this investigation that, with data of the type examined here, unbiased estimates of 50% points can be obtained by eye. Such estimates are very close to those obtained by the method of maximum likelihood and differences between individuals are likely to be very small indeed so that eye estimation provides a satisfactory substitute for the maximum likelihood approach. Fitting a straight line by least squares is slightly easier and quicker than a complete maximum likelihood approach but this method cannot really be preferred to eye estimation and is not a completely satisfactory substitute for maximum likelihood fitting of either $f_L(1|\underline{\theta})$ or $f_G(1|\underline{\theta})$. The moving-average method is not to be recommended generally. It can, even with moderately good data, give occasional rise to ambiguity by providing more than one 50% point for a given set of points. It only estimates the 50% point without bias and the calculation of the variance of the estimated 50% point is tedious.

A further interesting feature of the eye method is brought out in Table B-5 which gives the mean squared deviations of the eye-fitted curves, *i.e.* the values of n $\sum_{k=1}^{n} (p_j(1_k) - f^*(1_k))^2/n,$

 $f^*(1_k)$ being the eye-estimated proportion at length 1_k . It will be seen that differences between the eye-fitted curves are not too large, indicating that all seven people drew fairly similar curves over the whole range. These mean squared deviations are also similar to the same quantities calculated for the logistic curves fitted by maximum likelihood. Despite this the selection ranges estimated by the two methods do show consistent differences, the maximum likelihood estimates being larger in seven out of the nine cases (see Table B-6). The variances of the maximum likelihood estimates were obtained using

$$Var (1_1 - 1_2) = (1_1 - 1_2)^2 / b^2 S(nw1^2)$$
(2.3)

and are shown in Table B-6. The variances of the average eye-fitted values, also shown in Table B-6, were obtained in the usual way from the seven estimates for each haul. The variances of the eye-estimated selection ranges are more homogeneous than those of the maximum likelihood estimates, but both methods give similar percentage standard errors of around 10%.

This study has shown that the percentage standard error in the 50% retention length (or selection factor) estimated by somewhat elaborate statistical techniques is of the order of 1-2%, while that of the selection range is of the order of 10%. This sort of accuracy can be obtained, in general, without recourse to statistical procedures, simply by drawing free-hand curves by eye, although there is evidence that selection ranges estimated from eye-fitted curves may be smaller than by other methods, particularly for curves $f_L(1|\underline{\theta})$ and $f_G(1|\underline{\theta})$.

These results give some guidance in answering the problem posed at the beginning of this section, namely, how good are the basic data? Any set of data giving a percentage standard error in an estimate of the 50% retention length much in excess of 2%, say $\ge 5\%$, may be taken as unreliable. Any data suspected of being unreliable would have to have the accuracy of any estimate drawn from it estimated, and this could be done by either submitting the data to a complete statistical treatment or by having several people independently fitting curves by eye.

The studies discussed here were based on relatively good data, and the number of fish in each haul in the estimated selection ranges was in most cases quite large. For example, in the data corresponding to Tables B-3 and B-4, the numbers in the selection range varied between 273 and 4,160 fish. The variances of the maximum likelihood estimates quoted in Table B-4 are almost perfectly correlated with the number of fish in the selection range, but, even for the haul with the smallest number of fish, the percentage standard errors of the 50% point and the selection range are 1.3 and 13.5% respectively, both close to the average values quoted above.

In the next section it is pointed out that haul-to-haul variability is a much larger component of error than within-haul variability so that errors of estimation for a single haul are relatively unimportant in comparison with other errors.

3. <u>Haul-to-Haul Variability Within Experiments</u>

It is customary to make replicate hauls for selectivity study on the same ground and within a short space of time. Such replication is necessary if data are to be useful. The hauls to which the data in Tables B-1, B-2, B-3, and B-4 refer, for instance, were made within a space of three weeks, in both cases on the same fishing ground. They indicate the sort of variability encountered in such replicate hauls.

With $p_j(1_k)$ and $f_j(1_k)$ defined as at equation (2.1) we may write for the jth haul

$$p_{j}(1_{k}) = f_{j}(1_{k}) + \varepsilon_{jk}$$

$$= f(1_{k}) + \left\{ f_{j}(1_{k}) - f(1_{k}) \right\} + \varepsilon_{jk}$$

$$= f(1_{k}) + \beta_{j} + \varepsilon_{jk}$$
(3.1)

where $f(l_k)$ denotes the "true" selection curve for the conditions of the experiment and $\beta_j = f_j - f$ is the deviation of the expected curve for haul j from $f(l_k)$. The random term β_j is the haul-to-haul error. An equation, similar to (3.1), may be written in terms of any parameters estimated from individual haul data. Thus, without introducing any modification to the notation we might write

$$1_{j}^{(50)} = \lambda(50) + \beta_{j} + \epsilon_{je}$$
 (3.1)

where ϵ_{je} represents the error of estimation in the fitted 50% point for the jth haul and β_{j} represents, as before, haul-to-haul variability in the 50% points. If $\sigma_{\epsilon j}^{2}$ and σ_{β}^{2} are used to denote the true variances of the errors ϵ_{je} and β respectively, then

$$Var(1_{1}(50)) = \sigma_{\beta}^{2} + \sigma_{\epsilon_{1}}^{2}$$
 (3.2)

and, if $1_1(50)$, $1_2(50)$, ..., $1_n(50)$ are the estimated 50% lengths from n replicate hauls, on average the following equation is true,

$$\frac{1}{n-1} \sum_{j=1}^{n} \left\{ 1_{j} (50) - \overline{1}(50) \right\}^{2} = \sigma_{\beta}^{2} + \left(\sum_{j=1}^{n} \sigma_{\varepsilon_{j}}^{2} \right) / n$$

$$= \sigma_{\beta}^{2} + \overline{\sigma_{\varepsilon}}^{2}$$
(3.3)

where $\bar{\sigma}_{\epsilon}^{\ 2}$ is the average variance of the errors of estimate for each haul. Using the maximum likelihood estimates of the 50% points given in Tables B-1 and B-3, we find that the value of the quantity on the left-hand side of (3.3) is 14.2407 and 2.4225 respectively. The estimates of $\sigma_{\beta}^{\ 2}$ and $\bar{\sigma}_{\epsilon}^{\ 2}$ for Sets 1 and 2 are therefore

$$s_{\beta}^{2} = 14.1384$$
 $s_{\beta}^{2} = 2.3257$
 $\overline{s}_{\epsilon}^{2} = 0.1023$ $\overline{s}_{\epsilon}^{2} = 0.0968$

respectively, the quantities s_{ϵ}^{-2} being calculated from the variances given in the first columns of Tables B-2 and B-4. The variance of the between-haul component within an experiment is thus seen to be considerably larger than the within-haul component and, clearly, errors of estimation of the 50% point are insignificant. As pointed out by Pope (1966), in such situations, the most appropriate way to combine estimates from replicate observations is to employ a simple unweighted mean.

The value of $s_{\beta}^{\ 2}$ for the two sets treated here are quite different and reflect the much greater variability in the 50% points quoted in Table B-1 than in those quoted in Table B-3. In set 1 there is one quite low value (namely 18.2) and three relatively high values (namely 28.2, 31.2, and 32.9). Naturally the question arises whether or not one or more of these estimates should be rejected before calculating the experimental mean.

There are many possible reasons for encountering outlying observations. Selection may be reduced, for example, by quantities of weed blocking the codend meshes, etc. On the other hand the 50% point may be raised because of tears or large holes in the codend meshes, etc. It is clearly necessary in selection experiments to keep accurate records at the time, to which reference may be subsequently made for the possible elucidation of abnormal results. Of course this procedure is not without its pitfalls as it is deceptively easy to find, after an outlying observation has been observed, plausible reasons for its occurrence. Probably the most common situation is that in which there is no real evidence of a causal mechanism for the occurrence of a divergent observation and in such situations the idea of a test procedure for deciding whether to reject an observation or not has some apparent attractions. However, uncritical use of any such procedure is clearly dangerous. It is always advisable, wherever possible, to carry out any projected analysis of data both with and without the outlying observations. If the general conclusions are different in the two analyses then clearly the outlying observations are in a position to sway action one way or another. In such a situation no firm conclusion is warranted.

Reference has already been made to the extreme observations of Set 1 given in Table B-1. The unweighted mean and standard deviation of all 15 values of the 50% point are 25.3 and \pm 3.77 respectively. The observation with the maximum deviation is 32.9 and the deviation (+7.6) is twice the observed standard deviation. This is in no sense a large deviation, but, for interest, the mean and standard deviation have been recalculated omitting this observation. They are found to be 24.7 and \pm 3.25 respectively, not very different from the previous values.

The percentage standard deviations for the two sets of data considered here are 14.9% and 5.2% for Set 1 and Set 2 respectively. The latter figure is close to values quoted by Gulland (1964) for three sets of data obtained by the English research vessel Sir Lancelot, namely 5.2, 5.3, and 7.3%, and also to a value of 6.1% obtained by the author from a set of 20 replicate hauls. The average of these figures is 7.3% which may be taken as the percentage standard deviation of a single 50% point or selection factor within an experiment. The percentage standard error of the mean of a set of n replicate hauls (i.e. hauls within an experiment on the same ground) is then of the order of $(7.3/\sqrt{n})$ %. The actual standard errors of the mean 50% points for Sets 1 and 2 (data of Tables B-1, B-2, B-3, and B-4) are \pm 0.973 and \pm 0.519 respectively. These give rise to 95% confidence limits of 23.2 to 27.4 for Set 1 and 28.2 to 30.6 for Set 2. In terms of selection factors these limits are 3.31 to 3.91 for Set 1 and 3.48 to 3.78 for Set 2.

4. Variability Between Experiments

When several experiments have been conducted on the same codend material, the mathematical model given by equation (3.1) may be extended to include a component accounting for differences between experiments. Thus we may write

$$p_{ij}(1_k) = f(1_k) + \left\{ f_i(1_k) - f(1_k) \right\} + \left\{ f_{ij}(1_k) - f_i(1_k) \right\} + \epsilon_{ijk} = f(1_k) + \alpha_i + \beta_{ij} + \epsilon_{ijk}$$
(4.1)

where now $f(1_k)$ represents the true "global" selection curve for the codend material, $f_i(1_k)$ represents the true selection curve for the i^{th} experiment, $f_{ij}(1_k)$ represents the true curve for the j^{th} haul of the i^{th} experiment and ϵ_{ijk} is the true deviation of $p_{ij}(1_k)$ from $f_{ij}(1_k)$.

The component α_i may be a function of a large number of factors, e.g. season of year in which experiment was conducted, particular research vessel employed, experimental techniques such as duration of hauls, rigging of codend covers, etc. Gulland (1964) in a study of North Sea whiting selection factors for manila and sisal codends found significant differences between results reported by different scientists. Quoting Gulland, this "is not very surprising, as data presented by the same author are likely to be derived from observations on the same ground as well as with much the same gear". As an example of the sort of variation encountered in different experiments, the data given in Table B-7, taken from Pope and Hall (1964), are considered. These data refer to experiments using double polypropylene codends and the selection factors are for haddock.

The final column in Table B-7 gives the actual coefficients of variation for each separate set of results and they average about 5%. This agrees well with the between-haul within-experiment figure given in the previous section.

From equation (4.1), expressed in terms of selection factors x_{ij} instead of proportions $p_{ij}(1_k)$, the variance of a randomly selected value is given by:

$$Var(x_{ij}) = \sigma_{\alpha}^{2} + \sigma_{\beta}^{2} + \sigma_{\epsilon}^{2} \text{ or } \sigma_{\alpha}^{2} + \sigma_{\beta}^{2}$$
(4.2)

 (σ_{ϵ}^2) being small, as we have already seen). The variance of the unweighted mean selection factor of the ith experiment is:

$$Var(\overline{x}_i) = \sigma_{\alpha}^2 + \sigma_{\beta}^2/n_i$$
 (4.3)

where n_i is the number of replicate hauls in the ith experiment. For the data given in Table B-7, the quantities σ_{α}^2 and σ_{β}^2 may be estimated from a one-way analysis of variance. This is shown below, the quantity n_0 being $\left\{ \Sigma n_i - \Sigma n_i^2 / \Sigma n_i \right\} / 11 = 3.53$

	df	S. of S.	m.s.	E(m.s.)
Between experiments	11	1.22	0.1109	$\sigma_{\beta}^{2} + n_{o}\sigma_{\alpha}^{2}$
Within experiments Total	<u>41</u> 52	$\frac{1.14}{2.36}$	0.0278	σ _β ²

The estimates of $\sigma_{\beta}^{\ 2}$ and $\sigma_{\alpha}^{\ 2}$ are 0.0278 and 0.0235 respectively. Thus the component of variance between experiments is comparable with that within experiments. From these figures the variances of the experiment means are estimates as:

Experiment	$Var(\overline{x}_i)$
	$(0.0235 + 0.0278/n_{i})$
1.1	0.0235 + 0.0139 = 0.0374
1.2	+ 0.0139 = 0.0374
2.1 2.2	+ 0.0139 = 0.0374
· -	+ 0.0070 = 0.0305
3.1	+ 0.0139 = 0.0374
4.1	" $+ 0.0139 = 0.0374$
5.1	" $+ 0.0278 = 0.0513$
6.1	" $+ 0.0093 = 0.0328$
7.1	" $+ 0.0139 = 0.0374$
8.1	" $+ 0.0011 = 0.0246$
9.1	" + 0.0070 = 0.0305
9.2	" + 0.0093 = 0.0328

Since the values of n_1 are, with one exception, much the same in each experiment, the values of $Var(x_1)$ are very nearly equal. Hence a satisfactory mean for double polypropylene based on the 12 sets of data discussed here is an ordinary unweighted mean. In this case, the mean selection factor

is 3.47 with a variance given by $\left\{ \text{Var}(\overline{x}_1) + \text{Var}(\overline{x}_2) + \ldots + \text{Var}(\overline{x}_{12}) \right\} / 12^2 = (0.0235 + 0.0121) / 12 = 0.002967$. The square root of this quantity is \pm 0.0545. Thus approximate 95% confidence limits for the mean selection factor are 3.47 \pm 0.11 = 3.36 and 3.58. The two variance components, 0.0278 and 0.0235, correspond to percentage standard deviations of about 5%, thus confirming the magnitudes quoted in Section 3.

When combining the results of several experiments, especially in those cases where individual haul values are not available, a sufficiently reliable weighting system may be derived by assuming average percentage standard deviations for the within-experiment and the between-experiment components of 7% in both cases. For example, the ICES Cooperative Research Report, No.2 (1964) quotes an experimental value obtained by G. Saetersdal of 4.2, based on 11 hauls, for the selection factor of cod by a double nylon codend of 140-mm mean mesh size. The variance of this value may be taken as being approximately $(0.294)^2 + (0.294)^2/11 = 0.0943$. This method has been applied to data compiled by Treschev (see Part C) for double polyamide codends. Treschev divided published figures for this material into two groups which he referred to as polyamide A and polyamide B. The values of the cod selection factors and their estimated variances are shown in Table B-8. The latter will be seen to be fairly uniform, largely because the number of hauls in each experiment was reasonably large and similar from experiment to experiment. The inverses of the variances are quoted in Table B-8 and the weighted mean selection factors $(\Sigma_{\mathbf{W}_1}\mathbf{x}_1/\Sigma_{\mathbf{W}_1})$ for polyamides A and B are respectively 3.90 and 3.57. The weighted mean for both sets of data combined is 3.83. The variance of such a type of weighted mean is $1/\Sigma_{\mathbf{W}}$ and so the standard errors of these values are roughly \pm 0.051 and \pm 0.097 for polyamides A and B respectively and \pm 0.045 for both combined. This indicates approximate 95% confidence limits of 3.83 \pm 0.09 = 3.74 and 3.92 for double polyamide codends.

An additional form of weighting, not related to the accuracy of determinations of selection factor, may be deemed necessary when combining results derived from different experiments. It was pointed out that α_i may be a function of a number of factors operating during any experiment. Thus, for example, suppose it were known that the selection factor for a particular species was definitely lower in spring (say during February, March and April) than at any other time of year. If the majority of experimental results available referred to the spring period it might be more appropriate to calculate, in the above manner, two weighted means, one from experiments conducted in the spring (x_1) say and one from the remaining experiments (x_2) say. If, for some purpose, an average figure for the year as a whole were required and if the fishing effort by the commercial fleets was uniform throughout the year, a possible weighted mean would be $0.25x_1 + 0.75x_2$. Similar considerations would apply if it were known that selection factor was related to size of catch.

5. Mesh Differentials

We have seen in the preceding sections that the two main components of variance in an experimental value of a 50% selection length, or a selection factor, are the between-haul within-experiment and the between-experiment components. The coefficients of variation of both components are of the order of 5-10%, say 7% on average. Thus, for a codend material with an average selection factor of 3.3, the variance of a single determination would be 0.1067. If we assume that experimental values are normally distributed and further that such values are representative of the actual factors operating in the commercial fleets then 95% of all hauls made in a given period of time would have actual selection factors lying approximately between 2.7 and 3.9. Furthermore, if the average selection factor for another codend material were 3.6, and if selection factors for this material were also mercial hauls by the second material would have actual selection factors for this material were also mercial hauls by the second material would have actual selection factors below 3.3, the mean of the first material. There will in fact be a considerable overlap of two distributions with means differing ton factors for two different materials are really different or not, on which basis present mesh differentials have been established. But the existence of such overlapping distributions should not be in mean selection factors for different materials. The problem of establishing the existence of differences with the data available to him at the time, confidence limits were wider than the actual differentials than in existence.

With the accumulation of more and more data, however, greater precision in mean values should be achieved. In the previous section the average cod selection factor for double polyamide was

given as 3.83 ± 0.09 . If the mean selection factor for another codend material were established with the same accuracy, then it would be deemed to differ significantly from 3.83 if it were some 5-6% higher or lower.

The establishment of mesh differentials does not rest solely on statistical considerations, of course. Questions of distinguishing between different materials, and law enforcement are also very relevant.

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Table B-1. Comparison of 50% points obtained by different methods of estimation.

Table B-2. Comparison of variances of 50% points obtained by different methods of estimation (Set 1).

Maximum likelihood	Linear regression	Moving averages	Maximum likelihood	Linear regression	Moving averages
24.6	24.5	24.5	0.0395	0.0757	0.0371
26.5	26.5	26.6	0.0530	0.0350	0.0645
25.5	25.6	25.7	0.2707	0.3922	0.0043
21.9	21.7	20.7	0.1377	0.9239	
23.0	23.0	23.0	0.0905	0.0182	0.6999
25.5	25.4	25.7	0.1810	0.3094	0.1713
18.2	18.2	18.2	0.0046		0.4199
21.7	21.7	21.6	0.0134	0.0038	0.0053
27.5	27.5	27.5	0.0134	0.0045	0.0161
22.0	22.1	21.8	· =	0.0471	0.0784
25.6	25.6	25.6	0.1423	0.0319	0.2515
24.7	24.4		0.0533	0.1166	0.0593
28.2	28.2	24.5	0.1469	0.5138	0.2356
31.2	31.1	28.5	0.0776	0.0708	0.0876
32.9	_	31.1	0.2154	0.0182	0.1970
	32.6	32.5	0.0453	0.0759	0.0305

Table B-3. Comparison of 50% points obtained by different methods of estimation (Set 2).

Maximum	Linear	Moving				Evo f	itting			
<u>likelihood</u>	regression	averages	(1)	(2)	(3)	(4)	(5)	(6)	771	
28.2	27.6	28.3	28.2	27.8	28.1	27.7			(7)	Avg
29.6	29.4	29.8	29.8	29.7			28.1	28.1	27.6	27.9
29.7	29.5	=		,	29.7	29.8	29.7	29.9	29.7	29.8
26.3	-	29.3	29.3	29.4	29.0	29.4	29.4	29.4	29.4	29.3
• •	26.6	26.2	26.5	26.3	26.6	26.4	26.4	26.2	26.4	26.4
29.0	28.5	29.2	29.2	29.1	29.2	29.2	29.3	29.0	,	
30.5	30.4	30.4	30.8	30.5					28.9	29.1
31.6	30.8	31.1			30.7	30.4	30.6	30.7	30.5	30.6
29.2	28.9		30.5	30.6	30.5	30.3	30.4	30.4	30.7	30.5
–	· ·	29.1	29.6	29.2	28.9	28.9	28.8	28.4	28.8	28.9
30.8	30.5	30.9	30.8	30.9	30.8	30.8	30.8	30.9	30.8	30.8

Table B-4. Comparison of variances of 50% points obtained by different methods of estimation (Set 2).

Maximum	Linear	Moving	Eye
likelihood	regression	averages	fitting
0.2715	0.1663	4.7599	0.0562
0.0072	0.0524	0.0124	0.0062
0.0234	0.0173	0.2567	0.0224
0.0234	0.0536	0.0602	0.0167
	0.0550	0.0225	0.0190
0.0211	0.0535	0.0301	0.0200
0.0329	0.1935	0.0942	0.0181
0.2500		0.0825	0.1395
0.0632	0.1604		
0.1521	0.0794	0.2495	0.0024

Table B-5. Mean squared deviations from eye-fitted and logistic selection curves.

							Maximum
(1)	(2)	(3)	(4)	(5)	(6)	(7)	likelihood
0.0110	0.0108	0.0058	0.0101	0.0079	0.0086	0.0080	0.0091
0.0005	0.0022	0.0002	0.0009	0.0008	0.0016	0.0010	0.0013
0.0008	0.0014	0.0013	0.0013	0.0098	0.0014	0.0009	0.0012
0.0017	0.0020	0.0023	0.0026	0.0013	0.0015	0.0024	0.0027
0.0032	0.0022	0.0009	0.0035	0.0016	0.0054	0.0018	0.0031
0.0018	0.0013	0.0011	0.0003	0.0012	0.0016	0.0009	0.0014
0.0057	0.0072	0.0053	0.0099	0.0062	0.0097	0.0063	0.0062
0.0091	0.0085	0.0066	0.0086	0.0075	0.0103	0.0090	0.0068
0.0019	0.0011	0.0015	0.0011	0.0008	0.0010	0.0012	0.0017

Table B-6. Comparison of selection ranges estimated from logistic curves and eye-fitted curves.

_	ic curve likelihood)	Eye-fitt	ed curve
S.R.	Variance	Mean S.R.	Variance
6.35	2.7217	4.34	0.2162
5.17	0.0455	4.86	0.0662
5.87	0.1273	5.96	0.1262
5.13	0.3647	4.30	0.0600
6.27	0.1800	5.26	0.1462
4.05	0.0890	4.06	0.3195
9.97	2.0325	6.07	0.1290
5.22	0.4380	4.31	0.1581
5.16	0.4841	4.90	0.1533

Table B-7. Haddock selection factors for double polypropylene codends from different experiments.

Exper- iment	Grounds		Select facto		Mean	Variance	Coefficient of Variation
1.1	Fair Isle Buchan Deeps	3.5 3.5	3.7 3.8		3.6 3.6	0.0200 0.0450	3.9 5.9
2.1	Buchan Deeps Auskerry	3.5 3.7 3.8	3.6 3.9	3.6	3.6 3.8	0.0050 0.0167	2.0
3.1	Orkney/Shetland	3.4	3.0		3.2	0.0800	8.8
4.1	Scalloway	3.0	3.0		3.0	_	-
5.1	Moray Firth	3.7			3.7	***	
6.1	Buchan Deeps	3.5	3.2	3.3	3.3	0.0234	4.6
7.1	Moray Firth	3.4	3.3		3.4	0.0050	2.1
8.1	Faroes	3.2 - 3.3 - 3.4 - 3.5 - 3.6 - 3.7 - 3.8 -	- 1 -10 - 4 - 4		3.5	0.0298	4.9
9.1	Auskerry	3.7	3.6	3.4			
9.2	Scalloway	3.3 3.5	3.3	3.2	3.5 3.3	0.0333 0.0234	5.2 4.6

Table B-8. Average cod selection factors, variances and weighting factors for double polyamide polyamide codends.

		(a)	Double Pol	yamide A	
S.F. (* ₁)	V(x ₁)	$w = 1/V(\overline{x}_1)$	S.F. (x)	$V(\overline{x}_{1})$	$w = 1/V(\overline{x}_1)$
4.0	0.0844	11.8	3.2	0.0753	13.3
4.2	0.0943	10.6	3.7	0.0894	11.2
4.4	0.1186	8.4	3.5	0.0720	13.9
4.2	0.0951	10.5	4.0	0.0941	10.6
4.4	0.1044	9.6	3.8	0.0849	11.8
4.3	0.0951	10.5	4.0	0.0941	10.6
4.1	0.0865	11.6	3.9	0.0894	11.2
4.2	0.0951	10.5	3.5	0.0675	14.8
4.3	0.0997	10.0	3.9	0.0870	11.5
3.8	0.0778	12.9	3.7	0.0745	13.4
4.4	0.1044	9.6	3.8	0.0825	12.1
4.4	0.1044	9.6	3.4	0.0629	15.9
4.0	0.0862	11.6	4.1	0.0855	11.7
4.2	0.0951	10.5	3.8	0.0729	13.7
4.0	0.0862	11.6	3.7	0.1006	9.9
3.6	0.0762	13.1	3.5	0.0686	
4.1	0.1098	9.1		0.0000	14.6
		(b) Do	uble Polya	mide B	
3.7	0.0805	12.4	3.6	0.1270	7.9
3.7	0.0783	12.8	3.51	0.0679	
3.6	0.0688	14.5	3.38	0.0770	14.7
3.8	0.1061	9.4	3.38	0.0700	14.3
3.5	0.1200	8.3	3.50	V+ V04U	11.9

PART C

COMPILATION OF SELECTIVITY DATA FOR COD, HADDOCK AND REDFISH

by

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Introduction

The Working Group decided at a preliminary meeting held at ICES in October 1968 to restrict its analyses to data obtained from bottom trawl selectivity experiments made in ICNAF Subareas 1, 2, and 3 and ICES Sub-area I and Divisions IIa, IIb and Va in which cod, haddock, and redfish were the species being studies. Tabulation of the data was also limited to experiments conducted with double-braided twines because few, if any, trawls used in the above-mentioned areas of the North Atlantic are made of single-braided twines. All the data collected are given in Tables C-1 to C-12.

Only results from experiments based on the covered codend method were used. Experiments in which a topside chafer of any design was used were not included in the analysis. Experiments in which the selection factor was given in brackets or followed by a question mark were included but not those in which the number of fish in the selection range formed a very small proportion of the total catch. In those experiments in which the average duration of tow was not given, it was assumed to be 60 min. In the three component method of analysis described below, the total number of fish caught of the species studied was used as a weighting factor if the number of fish in the selection range was not stated; otherwise the result was not used in any of the four methods of analysis. Insofar as possible, all experiments made on one ground at one time were combined, that is, single hauls were grouped. This gave the following number of experiments which could be used:

	Manila	PP	PE	PES	PA
Cod	69	27	6	2	33
Haddock	47	16	1	2	16
Redfish	25	0	0	0	12

The Working Group was unable to investigate the effect of trawl construction on selectivity because there were insufficient data.

The Three Component Method of Weighting

In determining the mean value, the "weight" of the values investigated is usually determined by their frequency of occurrence and the mean is obtained from the formula:

$$\overline{x} = \frac{\Sigma(x_{\underline{i}}m_{\underline{i}})}{\Sigma m_{\underline{i}}}$$
 (1)

where x, is the value investigated,

m, is frequency of occurrence (weight) of value.

From the mathematical (statistical) viewpoint, this method of obtaining the mean value is fully justified. However, a more penetrating insight into the essence of the averaged trawl selectivity data obtained from different experiments easily shows that the true "weight" of each value of the selection factor is determined not only by its frequency of occurrence but also by the conditions of the experiment under which it is obtained. The most important factors are the number of fish within the selection range, which is a factor characterizing in each experiment the relation between the fish and the gear, and the duration of the experiment, i.e. the number and the duration of hauls.

Proceeding from this, a method of combining experimental data on trawl selectivity has been developed in the USSR (A.I.Treschev, VNIRO Proceedings, v. LXI, 1966).

According to this method, the mean weighted value of the selection factor is obtained from the formula:

$$\overline{k}_{s} = \frac{\sum_{i=1}^{n} t_{i} N_{i} k_{i}}{\sum_{i=1}^{n} t_{i} N_{i}} = \frac{n_{1} t_{1} N_{1} k_{1} + n_{2} t_{2} N_{2} k_{2} + \dots + n_{n} t_{n} N_{n} k_{n}}{n_{1} t_{1} N_{1} + n_{2} t_{2} N_{2} + \dots + n_{n} t_{n} N_{n}}$$
(2)

where

is the mean weighted trawl selection factor;

 $n_1, n_2 \dots n_n$ are the number of hauls in the first, second and n^{th} experiments;

 $t_1, t_2 \dots t_n$ are the durations of haul in the first, second and n^{th} experiments;

 $^{N}_{1},~^{N}_{2}~\cdots~^{N}_{n}$ are the numbers of fish within the selection range in the first, second and nth experiments;

 k_1 , k_2 ... k_n are selection factors in the first, second and n^{th} experiments.

It follows from (2) that the mean selection factor is directly dependent on the number of hauls, the duration of hauls and the number of fish within the selection range and, through all specific selection factors $(k_1, k_2 \dots k_n)$, on all other factors involved.

Thus, in applying this formula the numerical values of the selection factors obtained by the experimenters are taken as they are and in so doing it is assumed that the value of each specific selection factor reflects the heterogeneity of the populations fished, the size of catch, the peculi-arities of the trawl design and the fishing process as well as the method of combining data from different hauls, the method of constructing graphs and calculating selection factors, and many other factors. As far as the reliability (weight) of each specific selection factor is concerned, it is assumed that in all experiments it is mainly dependent on the duration of the experiment (the number of hauls multiplied by their durations) and on the number of fish involved, i.e. the number of fish mean value of the selection factor.

Tables of Selectivity Data

The following notes explain the headings of some of the vertical columns in Tables C-1 to

- SOURCE refers to Bibliography, Part E;
- 2) Number of fish in selection range refers to the total number of fish of the species studied in the selection range from all hauls combined.

	t (kg) of	Bycatch Cod-Cover	-											-												
	weight	Fish 1- Cover																								
	Av.	Cod- end																								
	Total No. of fish	died							675	229	1 1		732	g	527	1598	_ &	578	595	399	533	656	242	106	539	898
	Tota	0 9							4069	4517	-525**	-206-	4394	8967	1926	2535	141	1225	3232	451	1084	1522	768	674	1107	1485
	of fish selec-	Cover	Boat	Boat 2	Boat	Boat 2	Boat 1	Boat 2	113 d	ş	161	97	353	920	9	966E	33	416.	420	187	234	417			354	412
į	No. o: in se	tion range Cod- Cove	5485	2111	856	6065	1932	416	701	7	120		340	746	215	367	42	352	350	195	202	492			370	485
Va.	TION	Range (um)	100	Ξ	E		=	:	62 d	•	80	20	g	105	, a	P 601	g		105	100	120				120	120
IIb.	SELECTION	Factor	4.1	=	4.4	E	4.3	F	3.3	2.6	3.6	3.8	4	3.5	3.6	3.2	3.8	3.4	3.7	4.0	3.6	3,5	3.7	4.0	3.7	3.7
ICES I	205	length (mm)	540		580	=	570	¥	334	173	395	415	380		515	355	530		470	445	405		330	348	385	395
-		mesh 1 size (mm)	001	130	130	- 9	901	130	102	99	109	109		13	144	112	777	144	127	110	П				108	109
1.2		Mesh gauge					-				ບ	Ú			Aber		Aber		CNAF 4kg	Aber-deen	=	r			=	-
Area: ICNAF	Speed	of tow (knots)							3.0-3.8	u	3.5	3.5			М	=				3	3.9	3.8			3.6	3.8
Are	ILS ,	Average Duration (min)	85		105	z.	. 08	ε			60	09	17		06	ړ.	96	8		100	90	Š	9		06	E
	HAULS	No. Av	12	=	£	=	91	E	56	- 19	1	9			7	67			12	н	٠	2		n		2
Double Manila	METHOD		Parallel Haula	=	=	=	=	н	Cover	=		=	*	=	#	*	=	=	ı	:	z	=		E	c)	e) :-
tel: Dout		Runnage (m/kg)									121	121	160	. E					160							
Materi	COD-END	R Tex (g/1000 m)																								
awl		Material Const. (g									Twisted Continuous	æ	n	=			<u> </u>		Double						Twisted	£
Bottom trawl		DATE	V 1954	į u	r	=	z	E	1954	1955		=	1	ħ.	XI 1956	1956	TTI 1957	TV 1957		VII 1958	VII 1958	=	VII 1958	VIII 1958	VII 1958 1	£
Gears	1	Depth (m)									P.						190- 240	$\overline{}$		180	180- 200	190- 200			110-	200-
9	LOCALITY	Division Depth	ICES	ŧ		E	1	=	ICNAF 31,0,P	ŧ	ICES II b	=	±	Þ	ICES	ICNAF 3L,0.P.			ICNAF	ICES 1		. 1	ICES	=	ICES 1	" 2
	SOURCE	-	1963	=		z	-	\$	1957 d 1963	E	1956	±	19632	1956.	1963		_		1958	1958 ъ	=	=	1960	=	1958 ъ	E
TABLE C-1.	AUTHOR		SAFTERSDAL	н	Ε	z	E	=	TEMPLEMAN	=	BEVERTON	z	BRANDT	ŧ	SAETERSDAL	TEMPLEMAN	SAETERSDAL	T	MESTORFF	SAETERSDAL	=	ī	NOSSNOT	=	SAETERSDAL	=

^aICNAF Chafer; b

According to this method, the mean weighted value of the selection factor is obtained from the formula:

$$\overline{k}_{s} = \frac{\sum_{i=1}^{n} t_{i} N_{i} k_{i}}{\sum_{i=1}^{n} t_{i} N_{i}} = \frac{n_{1} t_{1} N_{1} k_{1} + n_{2} t_{2} N_{2} k_{2} + \dots + n_{n} t_{n} N_{n} k_{n}}{n_{1} t_{1} N_{1} + n_{2} t_{2} N_{2} + \dots + n_{n} t_{n} N_{n}}$$
(2)

where

is the mean weighted trawl selection factor;

 $n_1, n_2 \dots n_n$ are the number of hauls in the first, second and n^{th} experiments;

 $t_1, t_2 \dots t_n$ are the durations of haul in the first, second and n^{th} experiments;

 $^{N}_{1},~^{N}_{2}~\dots~^{N}_{n}$ are the numbers of fish within the selection range in the first, second and $^{n}_{th}$ experiments;

 k_1 , k_2 ... k_n are selection factors in the first, second and n^{th} experiments.

It follows from (2) that the mean selection factor is directly dependent on the number of hauls, the duration of hauls and the number of fish within the selection range and, through all specific selection factors $(k_1, k_2 \ldots k_n)$, on all other factors involved.

Thus, in applying this formula the numerical values of the selection factors obtained by the experimenters are taken as they are and in so doing it is assumed that the value of each specific selection factor reflects the heterogeneity of the populations fished, the size of catch, the peculiarities of the trawl design and the fishing process as well as the method of combining data from different hauls, the method of constructing graphs and calculating selection factors, and many other factors. As far as the reliability (weight) of each specific selection factor is concerned, it is assumed that in all experiments it is mainly dependent on the duration of the experiment (the number of hauls multiplied by their durations) and on the number of fish involved, i.e. the number of fish mean value of the selection factor.

Tables of Selectivity Data

The following notes explain the headings of some of the vertical columns in Tables C-1 to

- SOURCE refers to Bibliography, Part E;
- 2) Number of fish in selection range refers to the total number of fish of the species studied in the selection range from all hauls combined.

	(kg) of	Bycatch Cod- Cover end						-									_					-				
	ight	ver													_									:		
	Av. w	Fish Cod-Co end																								\neg
	fish	1 1							675	229			73	1010	139	1598	90	578	595	399	533	959	242	901	539	898
	Total No. of fish	Cod- end							4069	4517	-525-	-206-	7367	8967	3591	2535	141	1225	3232	451	1084	1522	768	674	1107	1485
	of fish selec-	Cover	Boat	Boat 2	Boat 1	Boat 2	Boat 1	Boat 2	۳3 ا	~	761	97	353	959		399 d	33	416	420	187	234	417			354	412
	No. of in sel	Cod- Cove	5485		856	6065	1932	416	108 d	~	170	124	340	342	-512	367 d	42	352	350	195	202	492			370	485
Va.	TION	Range (um)	100	Ε	£	=	=	=	P 29	•	88	8	. g	Ę	100	109 d	8	160	105	100	120	140			120	120 4
, IIb, Va	SELECTION	Factor	1.4	ż	4.4	E	4,3	=	3.3	2.6	3.6	3.8	3.6	5	3.6	3.2	8	3.4	3.7	4.0	3.6	3.5	3.7	4.0	3.7	3.7
I SZOI	50%	됩	240	E	580	E	570	£	334	173	395	415	g	097	515	355	550	495	470	445	405	385	330	348	385	395
	Меап	size (mm)	100	130	130	9	700	9	102	99	109	109	112	133	164	112	144	144	127	110	111	112			108	109
F 1. 2		Mesh gauge									٥	٥			Aberdeen		Aber	=	ICNAF 4kg	Aber- deen	ŧ	=			=	=
Area: ICNAF	Speed	– ,							3.0-3.8	=	3.5	3.5			m	=		m		m	3.9	3.8			3.6	3.8
Are	LS	Average Duration (min)	85	ı	105	ı.	80				09	99	17		90	ير .	9	90		100	90	105	60		90	=
	HAULS	No. Av	12	=	Þ		or	=	26	19	н	- 9	Ф	, va	7	29		- 2	12		2	2			м	
Double Manila	METHOD		Parallel Hauls		=	=	"	11	Cover	Ε	=	z.		Ŧ	#	=	=				£	E	z	=	# #	ह्य इ
		Runnage (m/kg)									121	121	160	"					160							
Material:	COD-END	R Tex (g/1000 m)																						_		
rawl		Material Const. (Twisted Continuous	£	=	=					Double						Twisted	.
Bottom trawl		DATE	V 1954	: "	r	*	*	=	1954	1955	п		=	Þ	XI 1956	1956	TTT 1957	TV 1957	VIII 1957	VII 1958	VII 1958	#	VII 1958	VIII 1958	VII 1958	ŧ
Gear:	III	Depth (m)															190-	100- 290		180	180-	190- 200	ı		110- 200	200-
	LOCALITY	Division	ICES II b	ŧ	=	×	#	u	ICNAF 3L,0,P	=	q II Sadi	ıı	ų	=	ICES	ICNAF 3L,0,P.			ICNAF	ICES		=	ICES	a	ICES	=
	SOURCE		1963				11	ı	1957 d 1963	H	1956	=	19613	1956	1963	1956d 1963	1963	ε	1958	1958 b	н	E	1960		1958 b	=
TABLE C-14	AUTHOR		SAETERSDAL	ı	=	t	F	ŧ	TEMPLEMAN	Ξ	BEVERTON	=	BRANDT	=	SAETERSDAL	TEMPLEMAN	SAETERSDAL	±	MESTORFF	SAETERSDAL	E	E	JONSSON	F	SAETERSDAL	_

^aICNAF Chafer;^b Given as 22 in Brandt,1956; ^CFixed Load Direct Pull Gauge 3 kg; ^dEstimated From Published Graphs and Tables.

	Av. weight (kg) of	Bycat	Cover Cod- Cover																								
	Aw.	Fish	-bog-										<u> </u>														
	Total No.	studied		9	+	-	+	+						'		1		<u> </u>	_	4059	 		900	5			2955
	Ĺ.,	\dashv	Cod-	5	3650	20 5	38			_	_	_	_	3877	- 7	5	16656	4133	3061	1805		2996	2,603	Ę	۶	,	2314
	of fish selec-	tion range	Cover	7,7	¥ 5	<u> </u>	1 1		1	<u> </u>	<u> </u>	<u> </u>	↓_		<u> </u>		<u></u> '	452	504	1236	2061	1197	8	5	Ę	197	1504
	No.		Cod-	─ ``	_	_			3					1300			7808	452	492	1014		1985	807.	137	7.67	130	1275
	SELECTION	г Капре	<u>(</u>	1	8		(18)		_	ļ		<u> </u>	ļ.	8	001)	88	8	8	110	120	160	140	02	68	(120	96	80
	SEL	1 12		-	3.4	,	, ,			;	3.3	,	,	3.8	(6.4)	4.0	,	3 4	7 6	3.5	3.0	2.8	2.5	3.4	3.1	3.5	3.1
		50% Length		09,	412	2	£ 5	;	ş	;	454	:	1 3	384 %	(9/4)	548	ě	353	7.5.7	449	371	350					398
	Į.	mean mesh	Size	113	120	Ş	Ę	=	2	=	137	=	-	101			911	Ş	200	123	E		109	151	-	128	128
		Mesh	gange	Aber-deen	_	م																	Scot-		7		
		, 	tow (knots)	3.5	3.5				ľ									3.5	3 5								
	HAULS	Average	Duration (min)	06	90	=	90		:	r	ŧ	2	:	=	1	ı		8	00		=	30	72	8	09	60	09
		Š.		1	5		2		~	~	=	-	-	1	3	2	12	9	10	20	4	1	4	1	1	6	2
	METHOD	_		Cover	Cover	=	=		=	£	=	=		Alternate Hauls	**	¥	ŧ	Cover	E	z	=	=	E	=	ŧ	±	ε
		Runnase	(m/kg)		101	=	125		=	=	Ξ	E	_	=						125	=	п	=	£	=	=	2
	COD-END	R Tex	(g/1000 m)									į															
		Material		Twisted	Twisted Continuous	t							<u></u>					Tylsted	=								
		DATE		VII 1958	VII 1959	=	VIII 1959	=	ŧ	:	Ŀ	E	=	E	E	=	ì	=	E	VIII 1959	=	=	:	:	Ξ	t	<u>.</u>
	Ħ	Depth	Œ	200-	Å Š	=									_		_	700	=	_			155				\dashv
	LUCALITY	Division Depth		ICES	ICES II b	Ξ	Ξ	z	=	=	-	=	F	=	=	=	=	=	*	ICES II b	=	E	ICES II b 1	:		E	
continued	SOURCE	_ 5_		1958 b	1959		1964	Ē	Ė	=	-		=	=	•	=	=	=	=	1964	=	=	=	=		-	=
TABLE C-1 (continued)	AUTHOR			SAETERSDAL	BEVERTON	=	ICES	£	r	Ė	E	ı.	14	11	=	=	E	, =	=	ICES		=	Ŀ		=	E	*

*ICNAF Chafer; Pixed Load Direct Pull Gauge 3kg.

l No. Av. weight (kg)	3 8		723	182	782	3563	1759	648	2005	1003	2711	7895	91	048	836	116	1890	921	009	259	702	12.5	830		- 00
Total No. of fish		306	٧,	535	345	1,70%	525	167						†							,			<u>.</u>	
of fish selec-	Cover	35	153	83	671	1,2%	182	313	567	777	307	8511					536			.,					
No. of In se	Cod- Cove	1.7	117	112	691	٠,											495 5				6 801				
HOIL	Range (mm)	(30)	S	(140)	2			85	82						160	180		100		,			٦	100	
SELECTION	Factor	3.8	7	3.6	3.5	3.5	3.7	1.7	3.2	3.2	3.6	α,	3.0	2 0	3.2	8	3.4	3.8	3.6	3.0	2.4	2.8	3.8		
ŧ	length	480	7.1	470	455	457	480	077	386	385	378	547	333	0%	369	324	447	494	470	703	254	296	689	513	
	mesh size (mm)	126			129					119							132 4	132 4	132 4	139 %				130	1
	Mesh gauge							8	8	ect							Aber- deen	ı	2	#	=	=	=	=	İ
**************************************	of of tow (knots)							3.5	3.5	3.5	3.5	ر د د					3.5	3.5	3.5	7	н	±.	£		
s	Average Duration (min)	80	2	90	60	80	09	06	50.	2				60.							5				
HAULS	No. Av	3	<u> </u>		3	3	9	9		2 45	10			4	09 60		2 80	2 90	4 90	7	7 10	90	90	90	ł
METHOD		er						er c													۳ م	q ·	Ü	U	
Ή	8,0	Cover	t .	*	-	=	=	Cower	=		±	=	Cover	#	_ ±	=	= '	=	=	=	Cover	=	Cover	=	
	Runnage (m/kg)	125	×	#	*	±	±	101	×	ź	125	=	2	=	4	=	125	=	=	r	=	r	=	ŧ	
COD-END	R Tex (g/1000 m)			ļ				m																	
	Material Const.							Twisted Continuou	=		=	:					Twisted	#	ŧ	E		E	*	:	
	DATE	VIII 1959	z	r	±	=	=	VII 1959	z	¥	VI 1960	ż	VII 1960	#	+1	=	VIII 1960	u	н	r		t			
III	Depth (II)							9 2			200						148-	\vdash	160	100-	160-	140- 200	120- 200	100-	I
LOCALITY	Division	ICES	=		=	ŧ		ICES II b	Ξ	F	ICES		ICES	ŧ	11	. =	ICES		=	Ξ	=		=	t	l
SOURCE		1964	-	=	=	=	E	(1959)	=		1964		1960	=		=	1960	z	22	=	ı		=		
AUTHOR		ICES	ı.	=	=	*	£	BEVERTON			ICES		TONSSON	ı	=		SAETERSDAL	=	E	t		2	=		-

AUTHOR	SOURCE	LOCALITY	III			COD-END		METHOD	≝	HAULS		 			SELECTION	<u>z</u>	No. of fish	E	Total No.	Av.	. Weight	(kg)	,
		Division	Depth (m)	DATE	Material Const.	R Tex (g/1000 m)	Runnage (m/kg)		No.	Average Duration (min)	Speed of tow (knots)	Mesh m	Mean mesh le size (mm)	S02 length (mm)	Factor R	T # 3	Cod- Cover	- J	studied od-	Cod	1 1/4	,L	
ICES	1965	ICES V &	191	VI 1962	Twisted	6623	151	Cover	80	8	0.4	ICES 1	127 4	401	3.2	3	330 370	 	266	729	, <u>t</u>		
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Total Catch of all Species (Codend and Cover).

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	3) of	Bycatch Cod-Cover										ļ	ļ	ļ		 							_		
	Av. weight (kg)	L										ļ	ļ												
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	Av	Cod- end											<u>_</u>												
	Total No. of fish	died	1257	95	260	153	148	255	83	183															
	Tota	stu Cod- end	1533	981	363	154	211	372	265	1223								į					-		
	fish lec-	ange	734		89	34																	4		
	No. of fish in selec-	tion range Cod- Cove	908	\Box	79	38			34	9						 									
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	HAULS	Average Duration (min)	8	ĸ	120	88	120	240	180	120										!					
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•	METHOD		Cover	=		=	=	=	=	Cover a Chafer															
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		Runnage (m/kg)	125	=	=	=	=	=	=	=							<u> </u>								
	COD-END	R Tex (g/1000 m)	•																			:			
		Material Const.	Twisted	=	=	=		:	н	11								-							
	- 1	DATE	IV 1964		ž	±	F	ŧ	F	=			-				<u> </u>	_							
	걾	Depth (m)	230- 250	097	=	t	E	460- 480	-09 1		-	<u> </u>	 								_				
(§)	LOCALITY	Division	ICNAF 2	Ī	=	=	<u>-</u>		*	*						 	-			l				-	
(continued).	SOURCE	Ö	1965	1966 2	ı		:	ı.	E	1966 c		-												-	
TABEL C-1					_					-			-	-		 	<u> </u>				 		ļ		
TABEI	AUTHOR		CRATBERG	HOLDEN	=	Ξ	E	=	ε	ŧ		ļ													

a Double Codend Manila, 126-mm mesh

TABLE C-2	-		Gear:	<u> </u>	Bottom trawl	Spec	Species: Cod	Material: Double Polypropylene	houble	Polyprop	4 analy	I isar	ICNAF 1	1.3: ICES	ᆑ	IIb. Va		ļ					
AUTHOR	SOURCE	LOCALITY	LITY			COD-END		METHOD	HAI	HAULS	Speed			204	SELECTION		No. of fis in selec-	fish lec-	Total No. of fish		Av. weight	ht (kg)	jo (
		Division	Depth (m)	DATE	Material Const.	R Tex (g/1000 m)	Runnage (m/kg)		No. Av	Average Duration (min)		Mesh E	mesh le size (H C	Factor	Range (III)	*	-	studied Cod-	1 3	Fish - Cover		I #I -~
HOLDEN	1966 d	ICES	170	X 1963	Plaited		216.0	Cover	-	ş	4	10801	2 2	5	-	1 5	+-	+		+	_		<u> </u>
OLSEN	1966	ICES	260-		e				4	8			1	†	†	┿	+	1	T	-		\perp	_
BRATBERG	1965	ICNAF	140-	IV 1964	Braided	1 4500		=	1	80	3.0		٠,	+	+	+					_		
HOLDEN	1966 a	ICES II b	460-	VI 1964	Plaited		216.0	Cover	-		4	=	I .	╁	<u> </u>	+	A7 87	3	3642 2618		<u> </u>	<u> </u>	
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BOHL	1966 a 1967 e	ICES V a	80 100	VI 1964	Continuou	\$067	204		=	ļ .	E	-	 	 	†	†	١.	T			;		
XARCETIS	1965 b	ICES V a	120-	7			700	-		\vdash	2		<u> </u>	1		 		1	7.0	<u> </u>	35	355	160
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BOHL	1967 c	ICNAF I B	180 - 220	XI-XII 1965	Continuous 4905	4905	204	:		 	•		121	3 2		8 6			9700	ľ	<u> </u>		
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OLSEN	1967	ICES	220- 250	111	Braided	4500		=	<u> </u>		1 6	 	191	807		92 93	6 6		516 933	-	59	173	38
BOHL	1967 а	ICNAF 1 F	100-	X 1966	Monofilament 4800	ent 4800	208	=	4		4	=	a	†	, ,	_			١,	1 2	 	, i	
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OLSEN	1968	ICES	220- 250	III 1966	Braided	4510		Cover a	7			ICRS 141		<u> </u>	_			Ī		<u></u>			
OLSEN	1966	ICES	290					Cover b		, -			-					265	- "	-	ļ		
HYLEN	1967	ICES	960-	v 1966	Braided	4510		Cover a	=	<u> </u>		=			1	1		1	۔ ا	ļ 1	<u> </u>		
HOLDEN	6961	ICNAF	220-230	XI 1967	Braided Split-Fibre 6203	te 6203	1615	Cover	7	٠	4			1			I —					679	۶
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Topside Chafer of Double Mesh Size (Polish Chafer); Double Codend: Chafer made of Manila.

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1969 " " " " " " 1969 " " " " " " " " " " " " " " " " " "

Bouble Codend: Polyester, Double Braided, Plaited, R 5000 tex 112-mm Mesh, Fastened Fore and Aft and along Selvedges.

1867 1968 1972	TABLE C-3	SOURCE	Gear: Bottom traul	rton tr	awl	Special	Species:Cod	Mat	Material: Double Polyethylene	uble	Polyethyl	- 1	Area: ICNAF 1.3:	MAF 1.	- 1	ICES L.I.B		No. of fish	┡	Total No.	\vdash				
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Av. weight (kg) of	Bycatch Cod- Cover					 				<u> </u>	<u> </u>	ļ. 		_	 								
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Av.	Fish Cod-Co																			 	.		
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SOURCE		1958																					
AUTHOR	-	MESTORFF	SAETERSDAL	=			:																

TABLE C-5	1	Geer: Bottom trawl	OCTOR	Trust	Speciestord	Cod Mate	Tieli	Desible Polys	7	mide Type A	Area: ICNAF 1	AVICO	1	1	· ICES I III V	Ye.				j				
AUTHOR	SOURCE	LOCALITY	E I			GINDA-GOOD.		METHOD	4	HADES	7000		5	Š	SELECTION	LION	No. of in se	of fish	Total No of flah	No.	Av. w	weight	(kg) o	jo
		Division	Depth (B)	DATE	Material Const.	R Tex (g/1000 m)	Runage (a/kg)		No.	Average Duration (min) (Mesh Sauge		41	Factor	Ken (II)	Cod- Cove	Cover	studied Cod-	+	Cod-Co		Sycatch Cod-Cov	tch
SAETERSDAL	1958 g	ICES I	150 220	X 1957	Twisted		260	Cover	п	8	8	Aber	140	85	4.2	ğ	T	707		9702		-	+-]
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Modified ICMAF Chafer; Double Codend-Full Length; Double Codend-Half Length; Tight Top Chafer Covering Half Codend Same Mesh Size

2	t.

Cod- Cover

Codend

112

3.7

114.4 409

CES

Cover a Chafer Cover b

Twisted

1CMAF 60- VI 3 and 4 '360 1965

1967

STRZYZEWSKI

Av. weight (kg) of

Bycatch

Fish

No. of fish Total No.
in selection fish
tion range studied

Code Cover Code
end end

Mesh mesh length Factor Range to (mm) (mm) c

No. Average of M Duration tow 8 (min) (knots)

Runnage (m/kg)

Material R Tex R Const. (g/1000 m)

DATE

Division Depth (m)

LOCALITY

SOURCE

AUTHOR

HAULS

METHOD

COD-END

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BUCKI ET AL	1968	ICNAF 3 K	330- 370	X 1967	z.	4700	133	11	7	240		:	110.5 410	410	3.7	107	209	323	1097	375				
Gear: R	Rottom trawl		Spec	Species:Cod	2	Material: Double Polyamide Type B	ouble Pol	yamide Iyı	e B	Area IC	ICNAF 1:	ICES	ICES IIB.	Va.										
AUTHOR	SOURCE	LOCALITY	ALIT			COD-END		METHOD	莊	HAULS	94		Mean	20%	SELECTION		No. of in se	of fish selec-	Total No. of fish	No.	Av.	veight	(8%)	of
		Division	Depth (m)	h DATE	Material Const.	R Tex (g/1000 m)	Runnage (m/kg)		No.	Average Duration (min)		Mesh gauge		# _	Factor	Range (mm)	Cod- Cove	14	studied Cod-	1ed	Cod-	rg S	Eycatch Ccd- Cov	Cover
BRANDI	1956	ICES II b		VII-VII	r Plaited		210	Cover	2		2.4	PEDGE 1	107	400	3.7	80	719	841		1270.				
=		ш		*			n		9				132	490	3.7	115	703	514	3556	845				
ž	1963	n					380	±	13		æĦ	ICES ITPE	901	410	3,8		168	172	73%	2031				
NESTOREF	1958	ICNAF 1 F		VIII 1957	Double		210	±	7,		H	ICNAF 4 kg	. 82	520	4.0	150	1005	1968	2672	379				
Ξ	2	' ε		=	3		280	н	6				129	500	3.9	56	515 8	866	2150	219				
ICES	1964	ICES II b	95- 110	VIII 1959	Plaited		161	±	4	60 3.	5-4.0		122	539	4.4	02	251 2		962	1675				
=	=	z	90- 105	z	=		210	ŧ	- 7	99		7	707	390	3.8	5[198	186	1171	538				
:	11	Ε	100- 110	п			210	=	-	36		_ 7	102	355	3.5	(130) (748)	(748) ((979)	1659	800				
=		ž	100-	ı.	ı.		210	E	-	9		7	102	370	3.6	90	1005	166	1749	1472				
	1965	ICES V &	80- 110	VI 1962			302	=	۷	9	4.0 I	SECTION OF THE PROPERTY OF THE	89	(310)	(3.5)	(80) 2	20 4	07		1				
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Polish Type Chafer with Mesh Size X4 Codend Mesh Size; Polish Type Chafer with Mesh Size X2 Codend Mesh Size;
Spring Loaded Gauge Subsequently Calibrated to ICES; d ICNAF Chafer 127.5-mm Mesh;

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	of.	i ii	Cover	T	2		J g	971				5.7		
	(kg)	Bycatch	Cod-		340		#		1	076	١,		1	59,
	Av. weight (kg) of	ا ا	Cover		375		3	1600 410 435	*	170		170 275		2683 1744 510 160 265
	Av.	Fish	-pog	+	975		7	1,600	1	310	⊢-	280	+-	510
	No.	led		1	687		†	2053		816		1054	-	1744
	Total No.	studied	Cod-		808		1	3050		619	_	1386		2683
	fish lec-	ange	Cover Cod-	†	†	;	Ť	, 6						1132
	No. of fish in selec-	tion range		-	=		+	613		(1601 (293)		(313) (418)	┞	1089
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(pen)	LOCALITY	Division Depth		ICNAF 1 d	TCMAT	1.5		=	:		;	=	ICES	II b
TABLE C-5 (continued)	SOURCE			1967 d		1967 a	_	1967 h					1968	1969
TABLE C-	AUTHOR			BOHI		=		£		+	-	+		

GROUPED DATA ALREADY TABULATED BY SINGLE HAULS IN PRECEDING TARLE

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Polish Type Chafer with Mesh Size X2 Codend Mesh Size; DICNAF Chafer 127.5-mm Mesh.

Material: Double menila Area: ICMAF 2,3; ICMS I, IIb, Va.	END METHOD HAULS Speed Mean 50% SELECTION in selection Av. weight	Runage No. Average of Mesh mesh 1 Duration tow gauge size (min) (knots)	1105 2359	66 194 (2.9) 33 7 17 1770	60 3.5 8 3.4 80 284 275 10		1 60 1.5 109 355 3.3	ICES 113 340 3.0 60 571 297	" 67 3-3.8 112 222 2.9 105 1872 1636 10348 15903 136 373	Parallel 4ber-152 575 3.8 90 82 112	" 1 90 (3) 100		1 90 (3) 133	Cover 6 90 (3) 144 490 3.4 120 638 544 1014 748	" 6 20 Scot- 7 280 2.9 65 1487 956 2175 1068	" 6 20 97 264 2.7 50 598 308 1380 x19	" 8 20 97 306 7.1 63 1560 1159 2530 1337	" 5 96 3.9 deem 111 352 3.2 80 90 91 498 775	151 Alternate 19 60 3.5 ICMAR 71 38920 1685 2112		" 5 " " 94 286 3.0 11490 ¥ 862	1500	" " 94 256 2.7 31312 codend	20%	" " 5 " " " 104 338 3.3 7805 677 1185	" " 5 " " 104 314 3.0 13835 1112 1446
_	SEE	h Facto	3.1	(2.9)	3.4	3.3	3.3	3.0	2.9	3.8		3.6		3.4	2.9	2.7		3.2		3,3	3.0	3.2	777	- 1	33	30
CES I,			313	194	365	355	355	3	222	575				9	280	264	30	352		8	286	298	, <u>2</u>	367	338	314
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AUTHOR		ğ		=	_	2	7		=		<u>.</u>	ī	2	•	SAETERSDAL		1	GHEV LNOV	SAETERSDAL		=				E	
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Aultiflap Chafer; Double Codend 105-mm Mesh; Gauge Calibrated with ICES Gauge.

No. Average Total Tota	SOURCE LOCALITY	tocality	Hi				COD-END	-	METHOD	\$	HAULS			<u> </u>		SELECTION	2	No. of fish in selec-	-	Total No. of fish	Av.	veight	(kg)	of.
Atternate 6 60 3.5 TCMM 104 314 3.0 God- Cover Cod- Cod- Cover Cod- Cod- Cover Cod- Cod- Cod- Cover Cod- Cod- Cod- Cod- Cod- Cod- Cod- Cod-	COD-END DATE	TTY COD-END	TTY COD-END	000 ENB	000 CO	COD-END	- E				98	70					7	in sele Ion rar		fish udied	121	≱ I Æ İ		
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* Fixed Load Direct Pull Gauge 3 kg; b Flap Type Chafer;

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1965 V	▎ ▘ ▎▕▍▕▘▕▘▕		†		(min)			mesh afze (m)	length (==)	Pactor	Range (III)	-	ы	studied Cod-	1 0	Fish	
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" " 138 " " 150 " " 150 " " 150 " " 150 " " 150			151		┢	3	-		Ž ŝ			1	Τ,		1818	238	
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150 " 150 " 150 " 150 " 167 " " 167 " " " 166 1966 d I 256			5		┼	;	West	7	q] _	_					177	10I - +
" " 150 83- " " 147 " " " 178- 1966 d I 260		6135	163		16 83	9		137.7	£ 2	9 -	(100)					268	
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AUTEOR	SOURCE	LOCALITY	III			COD-END		METHOD	HAULS	• •		<u>, 4</u>			SELECTION	No.	of fish selec-	Tot	al No.	Av. w	veight	(kg) of
		Division	Depth (m)	DATE	Material Const.	R Tex (g/1000 m)	Runnage (m/kg)		No. Ave Dur	Average Duration to (min) (kn		Mesh mesh gauge size (mm)		gth Factor	or Range		Cod-Cover	E Sod		Fish Cod- end	li di	Bycatch Cod- Cover
MIS 10	1966	ICES	260- 265	11 1964				Cover	4 80-	80-120 3.0		ICES 145	5 497	3.4	101	96	2	-	730	-		-
POHI	1966 a	ICES	100	VI 1964	Continuou	4905	204		14 120				 		2	252	2			375		35
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MARCETTS	1965	2		£	£		007	Ľ	5 120			717		- 4		$\overline{}$						
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HYLEN	1969	SZOI	² , 00	1366	Split- fibre	4510		=	10 80	3.5		136			7	1014	849	2828 4	4037		;	
CHS1180AVO	е	=	248	XI 1966				=	7 75	3,	Z5 toft	Lowest toft 120.	2 450		- R	(162)	(107)	7.11	258			
		*		ž				E	6 75	3.25		127	094 6.	3.6	118	(198)	(1.56)	764 4	418			
UNPUBLISHED (LI)	æ	=							6 75	3,25		5018- 5018- 107.4	4 345	3.2	118	(41)	(42)	658 II	128			
AZS 10	1968	r.	220- 250	111		4510		Cover b	3 74	3.0	3.0-3.5 ICES				733	85	105	145 30	307		- -	
BTT EN	1967		980	ν 1966		4510		£	10.80		-		4	7	- 3	1274	1515		2358			
UNPUBLISHED	е	*	235	XI 1967				Cover	9 51		<u>-</u>	105	5 370		g	184	270	1450 4	450			
	,	2	ε	2				E.	12 60	=	-	104.7	7 380	3.6	105	138	178	1259 3	350			
																						
	i		ŀ													_						
Double	Double Codend: Manila Chafer;	illa Chafe	٥	olish Ty	ype Chafer	Polish Type Chafer : Mesh Size	e Twice	that of Con	Codend.													

GROUPED DATA ALREADY TABULATED BY SINGLE HAULS IN PRECEDING TABLE

TABLE C-7 (continued).

0 0	S Et	<u> </u>		9				ļ. <u>.</u>	<u> </u>	<u> </u>	<u> </u>													ļ	
(kg)	Bycatch Cod-Cove																	"							
Av. weight (kg) of	sh Cover										†· -		ļ·					1	 	 <u> </u>					
Av.	Fish Cod-Co		36,		<u> </u>		T				 	-			†			 	† ·· ·	 					
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Total No. of fish	Studi Cod-		1521 63	1	-	†-	1		 	 	-		-			-	-			 	ļ				
		-					+	-		\vdash	 	 			 								-		
No. of fish in selec-	ra	 	392	T		<u> </u>	 	ļ			ļ		<u> </u>							 					
0 t	End-	-	328	-	<u> </u>	-	 	-	 		-		-	_	<u> </u>	ļ <u>.</u> .	ļ		<u> </u>	 					
SELECTION	T Ran	ļ	5	_		<u> </u>	ļ	-		ļ_	ļ	_		ļ						 					
SEL	Facto		3.3	<u></u>																					
) §	mesh length Factor Range (um)		804														_								
X S S	mesh size (wm)		24.4											-											
	Mesh gauge		ICES 124.4			-								-											
Speed	of tow (knots)				!						\ <u></u>														
HAULS	Average Duration (min)																	-						•	
HA	No. A	-	4 60				 	 	-	-	<u> </u>												_		
МЕТНОВ			Cover																	 					
	Runnage (m/kg)		204					-										-		 					
COD-END	â							<u> </u>												 					
8	R Tex (8/1000	<u> </u>	4905																						
	Material Const.		Continuous Plaited																						
	DATE		VII 1964																						
Y.I.	Depth (m)		70																	 				\neg	
LOCALITY	Division		ICES V &							-	<u> </u>					<u></u>				 					
SOURCE			1967 e			<u> </u>					-					·- 				 					
ALTHOR			BOHL										· · · <u>-</u> -			•			,	 _					

e e	Cover						<u> </u>																
(kg)	Bycatch Cod-Cov																 		-				
Av. weight (kg) of	'er										_											_	
Av.	Cod-Cov																 						
No.	lied	631	282														 		***				
Total No. of fish																							
of fish selec-	Cover	245	(228) 633																				
No. of fish in selec-	Cod- Cove	185	8						·									·					
Ior	Range (mm)	8	120																				
SELECTION	Factor	3.2																					
20%	length Factor 1 (mm)	391	283																				
Mean	mesh lusize	121		 -		 							-				 			-			
	Mesh gauge	SZOI															 				'	:	
Peed	of tow (knots)	3.5				1																	
	age tion n) (k		3	<u> </u>						.	ļ												
HAULS	Average Duration (min)	S	8	 <u> </u>		_														<u> </u>		 	
L d	No.	- 51		 		<u> </u>	<u> </u>												<u> </u>	<u> </u>			
		Cover	Cover																				
ME	Runnage (m/kg)																						
COD-END	R Tex (g/1000 m)																						
	ial 1			 	-	╂	-		_		-						<u> </u>		<u> </u>		<u> </u>		
	Material Const.					<u> </u>									ļ 								
	DATE	111	111																				
rry	Depth (m)	155-	250																				
LOCALITY	Division Depth (m)	ICES	#								· · · -			-									
SOURCE	Ğ							 				<u> </u>	<u> </u>				 <u> </u>		<u> </u>	-			
	ļ	1962	1964	_	_		-	ļ		_	<u> </u>	ļ	ļ		ļ	 	 <u> </u>			_	_	 _	
ALTHOR		OLSEX	=								<u> </u>												

1	S) of	Cod- Cover																					-		
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) jon	10.72	Cover		۶	 	_	_	_	ļ		ļ	ļ	ļ	ļ	ļ <u>-</u>	<u></u>	<u> </u>			<u></u>		ļ 	ļ		
-		8 5	_	75	_	ļ	<u> </u>	ļ	ļ			<u> </u>	ļ	_				<u>.</u>							
Total No.	fish	<u> </u>	2	1 2	ļ				ļ		<u> </u>		ļ		ļ										
	4 4	1	3	3 2	<u> </u>	ļ								L.,											
No. of fish	in selec- tion range	Cover	5	5	1_																				
No.			۶	8	1																				
SET ECT TON		Range (III)	8	ŝ																					
	1	Factor Range	3.3	3.5										!											
	502	Length (mm)	350	372																					
ICES I.	Mean	size (mm)	107.7	106.4		 			 	<u> </u>											-			 	\dashv
	400%		Aber-							 												-			
Area:		ot tow (knots)	6	l																					
_		1ge :1on 1) (k		3		-		 														·			
Polyes		Average Duration (min)	-30	8																					
ouble D	_	NO.	**				_	-	-	ļ															
Material: Double Polyester			COVET	=																					
Mate		Runnage (#/kg)	136	136																					
Species: Haddock		R Tex (g/1000 m)													-										
ea: B	-	(g/								-															
Speci		Material Const.							_																
trawl	ר ארה האדה		VIII 1960																						
MECON ITY		Depth (m)	145- 180	145- 170																					
Gear: Bottom trawl		Division Depth (m)	ICES I																						
SOURCE			1960	E												,									
TABLE C-9, AUTHOR S			SAETERSDAL																						
EAT.			SAETE																<u></u>	l 					

TABLE C-104	ı	GAT: MOLLOW LIBY		ļ	Sherien: HS	наддоск	Material	Material: Double Polysmide Type A	Pergra	aide Type	- 1	13:	J S		Area: [CNAF 2 3 4: [CPS] Va	4								
AUTHOR	SOURCE	LOCALITY	Z.I.T.			COD-END		METHOD	, н	HAULS	ā		Меап	50%	SELECTION	LION	No. of fish	No. of fish in selec-	Total No.	No.	Av. we	ight (Av. weight (kg) of	-
		Division Death	Penth	DATE	Material	20.7.0	Diminage		No.	No. Average		Mesh	mesn	Length	mesh length Factor Range	Range	tion range	ange	studied	eq	Fish		Bycatch	<u> </u>
			(ii)			(8/1000 m)	(m/kg)			Duration (min)	tow (knots)	gauge size (mm)	size (mm)	(EB)		Î	Cod-	Cover	Cod-	-	Cod- Co.	Cover er	Cod- Cover	Li di
SAETERSDAI	1958 a	ICES	150-	X 1957	Twisted		260	1000	=	5	(E)	Aber-	071	240	8	S	- 2	, , ,	705	0671	 -			Т
TRESCHEV STEPANOV	1968	= :	150	1959	=		182	=	-						3.8	1				3 ,		-	-	T
"	" "		200	1960	=			:	-4	6	3.5		+						361 77		<u> </u>			T
ICES	1965	ICES	148	VI 1962	=	3311	302	=	<u>~</u>			TCES			ļ · · ·		ر ا			1012	ļ	-	-	Т
ICES	1965	ICES V a	300	1962	=		182	=	יט					ĺ					۔ ا	-	-	-		Τ
	ı	n	300	u			=	=	<u>س</u>	08	3.5			368	 -	 	1			2784		ļ		1
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n	ıí	н	300	=	u		=	=	m		3.5					†	1	[1	7				Τ
	ı	=	300	=	£		=	5					t —		<u> </u>	Ι.] .	1		<u> </u>		\vdash	T
TRESCHEV STEPANOV	1968	ICNAF 2	250	1963	2		=	=	60		3.5	<u> </u>	 	1		+		Г	Ι.	, m	-	-	 	1
STRZYZEWSKI 1967	1967	ICNAF IM-4W		1965	ŧ		:	#.	۰.				4.	ļ		16	-			- 5				T

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	_ _ _		011	1390 865	415	225	330	665
			145	1390	250 965	415	665	320
			1115	240	250	215	5	140
			280	700	700	695	390	515
	175	1145	1563 280	384	31.5 700	3962 2830 695	2312 1297 390	2916
	359 175	710	1577	517	415	3962	2312	5213
	127	242	475	182		1026	647	105 1635 1677 5213 2916 515 140 320 665
	211 127	211 242	465 475	230	166 167	940	514 647	1635
	06	70	75	707	88	76	74.	105
	3.1	3,4	3.6	3.4	3.6	3.5	3.2	3.1
, уз.	330	350 3.4	479		929	438 3.5	007	
ES II	_	707	131.5	132.4 456	132.7	124.1	24.2	123.9 387
a: IC	ICES 107						H	
Are								
Tyne B	-			\dashv				-
er:1.c			116	9	09	SS	57	28
7017	٧	- 1	80			7	1 6	80
Material: Double Polyerile Type B Area: ICES IID, Va.	Cover	=	=	=	=	=	Cover a	Cover Chafer
Material;	210	380	210	÷	=	200	=	¥
dock			4760			5000	=	
Sear: Bottom trawl Species: Haddock			Continuous	=	<u>.</u>	=	=	Ε
1 Spe	VII-VIII 1956	=	VI 1964 G	VII 1964	z	VI-VII 1964	:	E
m traw			80	70	=	70	-	t
T: Botto	ICES II b		ICES V a		-	ICES V &	=	-
Gea	1963		1967 e	=	=	1966 g	=	=
	BRANDI	12	BOHL	±		=	:	=

GROUPED DATA ALREADY TABULATED BY SINGLE HAULS IN PRECEDING TABLE

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VI.	
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ICES	
1967	
BOHL	

a Multi-flap Chafer; b LCNAF Chafer (German Version)

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	(kg) of	Bycatch Cod- Cover								-				45 +									700	& EstoT	330	
	Av. weight	Fish Code Cover e	9	270	380			350	720	 				615 260 18	i						416 137	674 177	552 309	350 245 85	335	185
	Total No.	tudied I-	8	1	1	I	j ;	1		1			1104	1261	1422	320		200	3495			2113	928	k320	2936	
	<u> -</u>	<u> </u>	500	9	0598			7606	20212				1587	3 2076				637	— 1800			00 2484	575	3901	(2702) 2692	
	No. of fish in selec-	Cod- Cove	-	ļ		83				495 557	Г		1076 882	1557 1168	1185 466	349 230	-	549 301	1520 1306	203 173		2280 2000	500 830	2683 3669	(2005)	- 54
XTV.	-	Range (IIII)	78	25	54	6		 	2		- 5	1	96	150 1		103	111	108 5	67 1	78 2	70 4	160 2	250 5	162 2	(148)	205
ICES I.IIA.VA.XIV	SELE	Factor	2.7	2.7	2.1	3.1	2 8	2.6	2.5		2.6	3.0	2.9	2.6	2.5	2.7	2.9	3.0	2.9	2.9	2.2	2.2	2.8	2.9	3.1	1 1
ICES		-	300	1	141		405	8 292		310	380	007 9	383	386	292	313	8 352	8 358	380			275	369	396	3 469	260
ICMAF 1.3:		Mesh mesh gauge size (mm)	100AF 106 7	" 99.1	4.99	Aber- deen 144		ICNAP 111		129	Aber- deen 144	├─ ─ -⊦	132	West-	116	116	119.8	119.8	130	130.3	ICES 112	" 127	" 132	Aber ^a deen 138.	149.	901 3401
Area: IC	-		3.8			(3)					(3) de		(3)	We We	3.2	. 2.	. =		=	E ,	4.0 IC	4.0)	4.0	0 6
	1	Average Duration (min) (k												4		3.							40			
Double Manila	HAULS	No. Ave	99	_	01	1 90		7	39 30			- 38	1 90	5 70	8 90	8 90	9 9	90	 	90		8 60	3 60	11 108	10 174	•
Material: Do	8	-	Cover		E	=	2	=	=	z.	ŧ	2-14/4	t	ŧ	# .	. =		=	=	E	=	=	=		=	=
Mater		Runnage (m/kg)	182		252			182	202	160		125	125	120	153		=	-	2	153	151	ŧ	ž	163	120	
Redf 1sh	COD-END	R Tex R (g/1000 m)							.``				,,	8300							6623			6135		
Species: Redfish		Material R Const. (g/	Twisted	77	=	=	=	=	:	E	-	=		=		=	=	=		= :	· -	=	=		.89	:
rawl		DATE Ma	954		V-VII 1955	11 1956	IV 1956	¥ <u>9</u> ¥6	X-XI	· · · · · · · · · · · · · · · · · · ·	XI 1956	VIII 1960		VII 1961	1961		-			E .	VI 1962	z		VII 1962.	ı	IX
Gear: Bottom trawl		Depth (m)	A .	VI 19	V.	11 061	200- IV 275 195		нн	,	XI 246 19	200- VJ	200	390- VJ		350		350	350	350		179	161	10		
Gear: 1	LOCALITY	Division D	ICNAF 3N.0	r	ICNAF 3N.O.P		1CES 2	NAF.	ICNAF	ICNAF		SES 3	. 2	ICES 3				-	<u>د</u>	ICES 3	ICES V a 1			1 1	3	ICES
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TABLE	AUTHOR		TEMPLEMAN	±	:	SAETERSDAL	=	TEMPLEMAN	ı	BRANDT	SAETERSDAL	4	ŧ	BOHT.	IRESHEV	TRESCHEV	ко дис неч	ŧ	ŧ	IRESCHEV	ICES	=	=	ı	-	

a Calibrated to ICES.

		tch Cover							Π		1	1		+			. 1							_
	(kg) of	Bycatch Cod- Cov] 	-	-	-	្រូ ភូ		-	1					+	1.0		25	. 70	5.0	25	15	*	_
	ghe (5		Catch	ļ	-	-	Total	ጉ—		1	-			50	2	225	450	475	275	007	450	9	- 9	ļ.
	Av. weighτ	Fish - Cover	1	-	<u> </u>	_	;		<u> </u>					180	. 225	150	135	i	215	225		250	541	!
	Av	Cod-	1,5	-	_	<u> </u>	1	-					L	450	630	495	979	1225	960	1340	1540	1440	2,40	!
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PART D

CONCLUSIONS AND RECOMMENDATIONS

Conclusions

The Working Group accepted that the minimum mesh size of codends made from different materials should be determined in accordance with the principle of equivalent selectivity.

The selectivity equivalent of a given material is defined here to be the ratio of the average selection factor for that material to the average selection factor for double manila. The basic data available to the Working Group for the calculation of average selection factors were those listed in Tables C-1 to C-12.

Although the Group agreed that more work is needed on this problem, it decided to separate the selection factors in the tables of selectivity results for cod and haddock into two groups, those in which the polyamide used had an elongation of greater than 25% and those which had a lower elongation. The former was called polyamide A and the latter polyamide B. The division was made on the basis of the results given in Part A of this report, and it was necessary to assume for this purpose that the polyamide used was produced in the country conducting the experiment. This breakdown is shown in the tables of selectivity data (Tables C-1 to C-12).

The Working Group spent some time considering the most appropriate method of evaluating the average for each material. Besides the methods described in Parts B and C of this report, two other types of average were considered, namely the unweighted mean and the mean weighted by the number of hauls corresponding to each selection factor. The Working Group decided to use all four methods, and accordingly four mean values were calculated for each material. These mean values are listed in Table D-1, where it will be seen that, in general, they are in close agreement.

The selectivity equivalents for each material, obtained from the results shown in Table D-1, are given in Table D-2.

Unweighted selection factors for the experiments used are given in the form of a frequency distribution in Table D-3.

The Working Group discussed the selection results in relation to the physical properties of the various twines. It considered that there might be a relationship between elongation and selection factor and it noted that in the case of polyamide there is a wide range in elongation as defined in Part I of this report.

Recommendations

The selectivity equivalents for mesh sizes of trawls made from different materials obtained as a result of the analysis of selectivity data and of the tests of the physical properties of trawl twines and mentioned in the conclusions are recommended to be introduced in the Conventions for the North-East and Northwest Atlantic Fisheries in place of the present rules for establishing mesh sizes of trawls made from synthetic materials, insofar as the Commissions consider these groupings enforceable.

The Working Group recommends that for all future selectivity experiments the following information about the codend material used should be tabulated:

TWI	NE DATA	
*a.	Material	Chemical nature
*b.	Type of fibre	Multi-filament, mono-filament, staple fibre, film (= split fibre)
, с.	Construction	Twisted, plaited
*d.	Rtex	
*e.	Runnage	(m/kg)
f.	Treatment	Untreated, thermo-fixed, chemically treated, etc.
g.	Wet flexibility	
*h.	Wet weaver-knot breaking load	
1.	Wet load elongation curve	(a) during breaking load test,(b) after some repeat loadings
*j.	Wet breaking lengths (based on dry material)	
NET	DATA	
a.	Method of manufacture	Hand or machine made (knotted,

c. Wet mesh size using ICES mesh gauge

d. Wet mesh breaking load

b. Treatment

e. Wet load elongation curve

(a) during breaking load test,

treated, etc.

knotless, single or double braided)

Untreated, thermo-fixed, chemically

(b) after some repeat loadings

These tests are proposed in the present state of knowledge. The list is not exclusive and should be supplemented by other tests as and when they become available.

Tests marked with an asterisk (*) should be carried out according to ISO standards.

The following laboratories are prepared to undertake these tests, free of charge, for those institutions which do not have the necessary facilities:

Institut für Fangtechnik, Hamburg (Germany), Palmaille 9 (marked to the attention of Prof. A. von Brandt);

Laboratorium voor Materiaal Onderzoek TNO, IJmuiden (marked to the attention of Dr J. Reuter).

A sample of 100-m twine (minimum) and a piece of netting measuring 20 meshes by 20 meshes should be sent. To facilitate receipt it is absolutely essential that the customs declaration should be marked "No Commercial Value".

As far as other data are concerned, the Working Group recommended that the modified ICNAF form of tabulation used in this report should be adopted.

Because manils is no longer used commercially and because it is not readily evailable, the Working Group recommends that a new standard material be introduced.

The standard netting yarn chosen is made of PA-fibres and is manufactured by Apeldoornse Nettenfabriek in Apeldoorn, Netherlands. It is a twisted trawl yarn with the designation 3/500. The following list indicates the properties:

Designation	R 6484 tex
Twist direction	Z
m/kg	154 m/kg (= 77 yards/1b)
Diameter	2.87 mm
Twist level	185 approximately
Breaking load (dry)	355 kp approximately
Breaking load (wet)	315 kp
Knot breaking load (wet)	299 kp
Breaking length (dry)	54.9 km
Breaking length (wet)	48.6 km
1/2 knot breaking length	23.1 km
Extension (1/2 knot breaking load, wet):	23.8%

The factory mentioned above offers this PA standard netting yarn for 7.75 hfl (cif) in bunches of 2.2 kg. Customers are asked to order quantities of not less than 50 kg.

To ensure the uniformity of the properties on the one hand and to avoid deterioration during storage on the other, the factory intends to produce the netting yarn in quantities of 300 kg at a time.

If the relevant committees of ICES and ICNAF agree to adopt the new standard net yarn, the Working Group recommends that the committees take steps to determine whether a standard netting can also be manufactured, and, if so, that it be introduced for selectivity experiments.

The Working Group recommends that further work is needed to determine the relationship between elongation and selection factors.

Table D-1. Mean selection factors.

			Cod				Had	Haddock				2	Redfish		
	(1)	(2)	9	(4)	(4) Mean	Ξ	(1) (2)	(3)	(†)	(3) (4) Mean (1) (2) (3) (4)	\Box	(2)	(3)	(4)	Mean
Double manila	3.52	3.52 3.48	3.45	3.45 3.42 3.47	3.47	3.26	3.16	3.26 3.16 3.09 3.22 3.18	3.22	3.18	2.71	2.62	2.64	2.71 2.62 2.64 2.62 2.65	2.65
Double polyethylene	3.43	3.43 3.42	3.37	3.41 3.41	3.41	3.20	3.20	3.20 3.20 3.20 3.20 3.20	3.20	3.20					
Double polypropylene	3.74	3.74 3.70	3.60	3.60 3.63 3.67	3.67	3.48	3,45	3.48 3.45 3.42 3.45 3.45	3.45	3.45	1	1	ı	1	1
Double polyester	3,95	3.95 3.95	3.95	3.95 3.95 3.95	3.95	3.40	3.40	3.40 3.40 3.42 3.39 3.40	3.39	3.40	ı	ı	I	ı	ŧ
Double polyamide (A + B) 3.92 4.00	3.92	4.00	4.02	3.89 3.96	3.96	3.59	3.66	3.66 3.62 3.63 3.63	3.63	3.63					
Double polyamide A	4.09	4.09 4.09	4.03	4.07 4.07	4.07	3.67	3.86	3.67 3.86 3.73 3.67 3.73	3.67	3.73	ı	ı	1	ı	1
Double polyamide B	3.66 3.66	3.66	3.57	3.57 3.63 3.63	3.63	3.40	3,33	3.40 3.33 3.51 3.40 3.41	3.40	3.41	2.92	2.95	2.84	2.92 2.95 2.84 2.87 2.90	2.90
				ļ											

(1) = unweighted mean;

(2) * weighted by number of hauls only;

(3) = weighted by method described in Part C;

(4) = weighted by method described in Part B.

Table D-2. Characteristics of different net materials and equivalents.

	Characteristics of material		E	quivalents	3
	Percentage elongation at the load of the half knot				
<u>Material</u>	breaking load		Cod	Haddock	Mean
Double manila	6.5- 7.8		1.00	1.00	_
Double polyethylene	8.4-22.2	(1)	0.97	0.98	
		(2)	0.98	1.01	
		(3)	0.98	1.04	_
		(4)	1.00	0.99	-
		Mean	0.98	1.00	0.99
Double polypropylene	10.4-21.3	(1)	1.06	1.07	_
		(2)	1.06	1.09	_
		(3)	1.04	1.11	-
·	· · · · · · · · · · · · · · · · · · ·	(4)	1.06	1.07	_
		Mean	1.06	1.08	1.07
Double polyester	8.3-12.3	(1)	1.12	1.04	
	•	(2)	1.14	1.08	-
•		(3)	1.14	1.11	-
		(4)	1.15	1.05	
		Mean	1.14	1.07	1.10
Double polyamide	15.5-47.0	(1)	1.11	1.10	_
•		(2)	1.15	1.16	-
		(3)	1.17	1.17	
		(4)	1.14	1.13	<u>-</u>
		Mean	1.14	1.14	1.14
Double polyamide A	>25.0	(1)	1.16	1.13	_
		(2)	1.18	1.22	-
		(3)	1.17	1.21	-
<u> </u>		(4)	1.19	1.14	
		Mean	1.18	1.18	1.18
Double polyamide B	≤ 25.0	(1)	1.04	1.04	-
		(2)	1.05	1.05	-
		(3)	1.03	1.14	-
		(4)	1.06	1.06	<u>-</u>
		Mean	1.04	1.07	1.06

^{(1) =} unweighted mean;

^{(2) =} weighted by number of hauls only;

^{(3) =} weighted by method described in Part C;

^{(4) =} weighted by method described in Part B.

Table D-3. Frequency distributions of unweighted selection factors for cod, haddock, and redfish.

		Haddock					Redfish					
S.F.	Manila	PP	PE	PES	PA	Manila	PP	PE	PES	PA	Manila	PA
2.1											1	
2.2											2	
2.3											0	
2.4											1	
2.5											3 3	
2.6												1
2.7 2.8	3					1					3	4
						0					2	1
2.9 3.0	3 2					4 5					5 2	2
											•	0
3.1 3.2	2 2	3	1			3 7	1	1		1 0	2 1	2
3.3	4							1	_		ī	0
3.4	15	5 0	1 2		1	11 6	2 6		1 0	0 2		1 0
3.5	7	4	1						1			
3.6	7	1	Ō		5 2	6 2	1 3		1	3 5		1
3.7	6	1	0		3	1	2			1		
3.8	8	ī	1		4	ī	ī			3		
3.9	0	0		1	2					0		
4.0	5	3		1	5					1		
4.1	3	3			2							
4.2	0	4			4							
4.3	2	1			2							
4.4		1			2 3							
Cotal	69	27	6	2	33	47	16	1	2	16	25	12

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