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Greenland Halibut SCAA Robustness Tests

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Abstract

This paper provides details of the full set of SCAA (Statistical Catch At Age) operating models intended for use in the MSE process for Greenland halibut. It first summarises the Reference Case operating model and the six variants thereof to serve as robustness tests, and then provides values of key associated parameter values and plots of estimated trends for various components of the biomass. Two Appendices list the data used for conditioning (fitting) these models and the complete mathematical specifications of the models and the (penalised) likelihood maximised in the fitting process.

Introduction

The Management Strategy Evaluation (MSE) process for Greenland halibut is to take account of two sets of Operating Models (OM) – the one set based on assessments conducted using XSA, and the other on SCAA-based assessments.

The following are the SCAA Reference Case (RC) and robustness test OM for this process:

- 1) Reference Case: Case 2 of Butterworth and Rademeyer (2010): Beverton-Holt, $h=0.9$, $M=0.2$, exponential decrease in selectivity for ages 11+;
- 2) RC with flat commercial selectivity (estimated) for ages 11+;
- 3) RC with flat commercial selectivity (fixed similar to XSA value) for ages 11+;
- 4) RC with $M=0.1$;
- 5) RC with $M=0.2$ for ages 0-10, linear increase to $M=0.4$ for age 14; and constant thereafter;
- 6) RC with $h = 0.6$ in the assessment, to simulate a stock that has a large maximum recruitment which has been severely recruitment-overfished;
- 7) RC with a modified Ricker stock-recruitment relationship: $R_y = \alpha B_y^{sp} \exp(-\beta(B_y^{sp})^r)$.

Data and Methodology

The data used to fit these models within an SCAA framework is given in Appendix A. Note the fecundity-at-age matrix in Table A.4 is not the most recent version of this matrix as advised by Healey (pers. commn). The results following are based on this earlier matrix. Re-runs of these SCAA cases using these updated data will be used for

the final MSE analyses. However the differences in their results from the plots and results given in the Tables and Figures below are minimal.

The SCAA methodology applied is detailed in Appendix B.

Results and Discussion

The results of the SCAA variants explored are listed in Table 1, with corresponding biomass trajectories plotted in Fig. 1 (which also includes the baseline XSA results for ease of comparison) and stock-recruitment relationships shown in Fig. 2. Table 2 contains the estimated numbers-at-age matrix for the Reference Case, and Table 3 lists the estimates required should another wish to project trajectories forward for any of these OMs.

The biomass trajectories in Fig. 1 show that the various robustness tests generally span the range between the SCAA and XSA Reference Cases, with differences in scale for the 10+ biomass arising primarily from the different ways in which the commercial selectivity for ages 11+ is set.

Table 1: Results of fits of SCAA Reference Case and the intended robustness test operating models (see text for details) to the commercial catch and survey data. Values fixed on input rather than estimated are shown in **bold**. Quantities shown in parenthesis are Hessian-based CVs.

	1) Reference Case	2) flat commercial selectivity for ages 11+ estimated			3) flat commercial selectivity for ages 11+ fixed to XSA value			4) $M = 0.1$			5) $M_{10}=0.2, M_{14}=0.4$, linear in between			6) $h=0.6$			7) Modified Ricker				
'-lnL:overall	-630.8	-625.3			-610.3			-631.3			-629.3			-629.5			-631.0				
'-lnL:Survey	-29.9	-30.2			-30.7			-27.0			-30.3			-32.3			-30.2				
'-lnL:CAA	-222.8	-220.6			-203.1			-221.8			-223.2			-221.8			-222.1				
'-lnL:CAAsurv	-462.8	-463.6			-463.4			-464.8			-461.7			-459.8			-462.5				
'-lnL:RecRes	17.6	17.6			23.1			15.9			18.7			17.4			17.0				
'-lnL:SelPen	67.0	71.5			63.9			66.3			67.2			67.1			66.7				
h	0.90	0.90			0.90			0.90			0.90			0.60			0.80	$\gamma = 0.015$			
M	0.20	0.20			0.20			0.10			0.2-0.4			0.20			0.20				
θ	0.31	0.18			0.04			0.28			0.35			0.59			0.31				
ϕ	0.28	0.34			0.46			0.29			0.27			0.13			0.26				
ρ - surveys	0.60	0.57			0.50			0.65			0.57			0.47			0.58				
ρ_{CAAage}	0.28	0.35	0.35	0.28	0.35	0.35	0.50	0.29	0.34	0.50	0.28	0.35	0.49	0.29	0.36	0.50	0.29	0.36	0.49	0.28	0.35
ρ_{CAAyr}	-0.32	-0.49	-0.49	-0.32	-0.49	-0.49	-0.69	-0.32	-0.49	-0.69	-0.32	-0.49	-0.69	-0.32	-0.49	-0.69	-0.32	-0.48	-0.69	-0.32	-0.49
K^{sp}	340 (0.06)			328 (0.06)			425 (0.06)			1056 (0.06)			181 (0.06)			420 (0.12)			374 (0.34)		
B^{sp}_{2008}	37 (0.42)			25 (0.58)			9 (0.27)			169 (0.33)			22 (0.46)			65 (0.45)			37 (2.33)		
B^{5-9}_{2008}	128			114			94			123			129			118			125		
B^{10+}_{2008}	53			38			16			198			39			85			53		
MSYL ^{sp}	0.18 (0.14)			0.19 (0.18)			0.20 (0.07)			0.20 (0.15)			0.17 (0.13)			0.33 (0.15)			0.16 (0.54)		
B^{sp}_{MSY}	60 (0.17)			61 (0.19)			83 (0.09)			206 (0.19)			31 (0.16)			138 (0.20)			60 (0.25)		
MSY	27 (0.05)			27 (0.06)			38 (0.06)			28 (0.06)			28 (0.05)			21 (0.12)			26 (0.42)		
σ_{comCAA}	0.07			0.07			0.08			0.07			0.07			0.07			0.07		
Survey	$q's \times 10^6$	σ_{surv}	$\sigma_{survCAA}$	$q's \times 10^6$	σ_{surv}	$\sigma_{survCAA}$	$q's \times 10^6$	σ_{surv}	$\sigma_{survCAA}$	$q's \times 10^6$	σ_{surv}	$\sigma_{survCAA}$	$q's \times 10^6$	σ_{surv}	$\sigma_{survCAA}$	$q's \times 10^6$	σ_{surv}	$\sigma_{survCAA}$	$q's \times 10^6$	σ_{surv}	$\sigma_{survCAA}$
CanFall	421	0.17	0.02	454	0.16	0.02	508	0.15	0.02	640	0.22	0.02	416	0.17	0.02	435	0.15	0.02	426	0.17	0.02
EU	218744	0.28	0.05	237401	0.28	0.05	264633	0.29	0.05	243516	0.28	0.05	209096	0.27	0.05	198912	0.26	0.05	218826	0.28	0.05
CanSpr	22	0.41	0.05	23	0.40	0.05	25	0.41	0.05	29	0.42	0.05	22	0.41	0.05	22	0.40	0.05	22	0.41	0.05
σ_R _out	0.21			0.21			0.24			0.20			0.22			0.21			0.21		

Table 3: Quantities needed for projections for the SCAA Reference Case.

h	0.900		
K^{sp}	339.94		
σ_R	0.210		
Survey biomass indices:			
	$q's \times 10^6$	σ_{surv}	λ_{2008}
CanFall	421.005	0.170	-0.017
EU	218744	0.281	-0.514
CanSpr	22.109	0.406	0.061
Selectivities			
age	CanFall	EU	CaSpr
0	0.285	0.050	0.068
1	0.533	0.063	0.159
2	1.000	0.079	0.375
3	0.696	0.126	0.623
4	0.769	0.178	0.728
5	0.664	0.491	1.000
6	0.515	0.965	0.658
7	0.463	1.000	0.370
8	0.279	0.800	0.127
9	0.116	0.369	0.044
10	0.060	0.283	0.015
11	0.041	0.155	0.005
12	0.027	0.085	0.002
13	0.018	0.047	0.001
14	0.012	0.026	0.000
15	0.008	0.014	0.000
16	0.006	0.008	0.000
17	0.004	0.004	0.000
18	0.003	0.002	0.000
19	0.002	0.001	0.000
20+	0.001	0.001	0.000
			0.031
			0.023

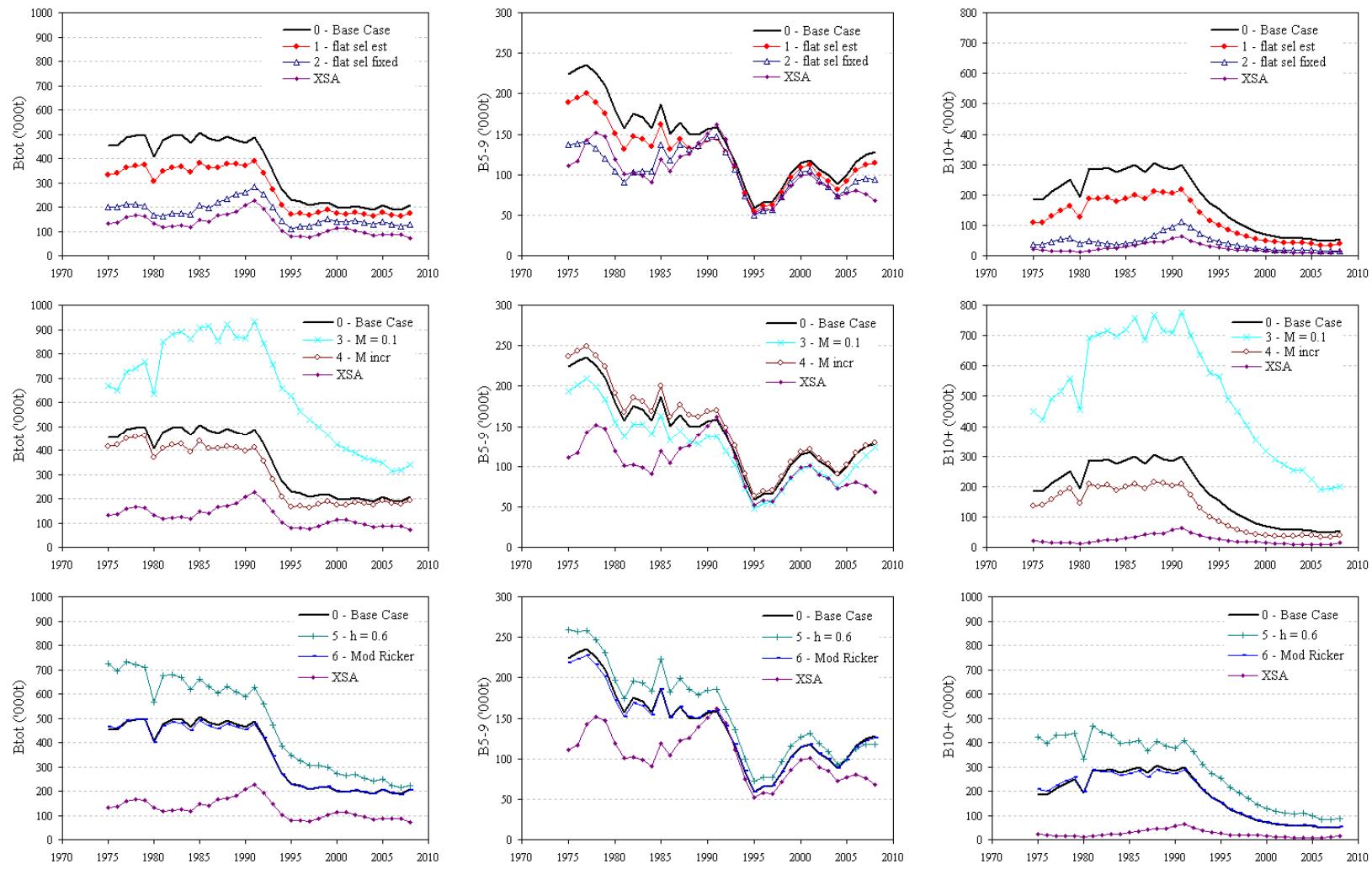


Fig. 1: Biomass trajectories for a series of SCAA variants and the 2008 XSA (Healy and Mahe, 2009).

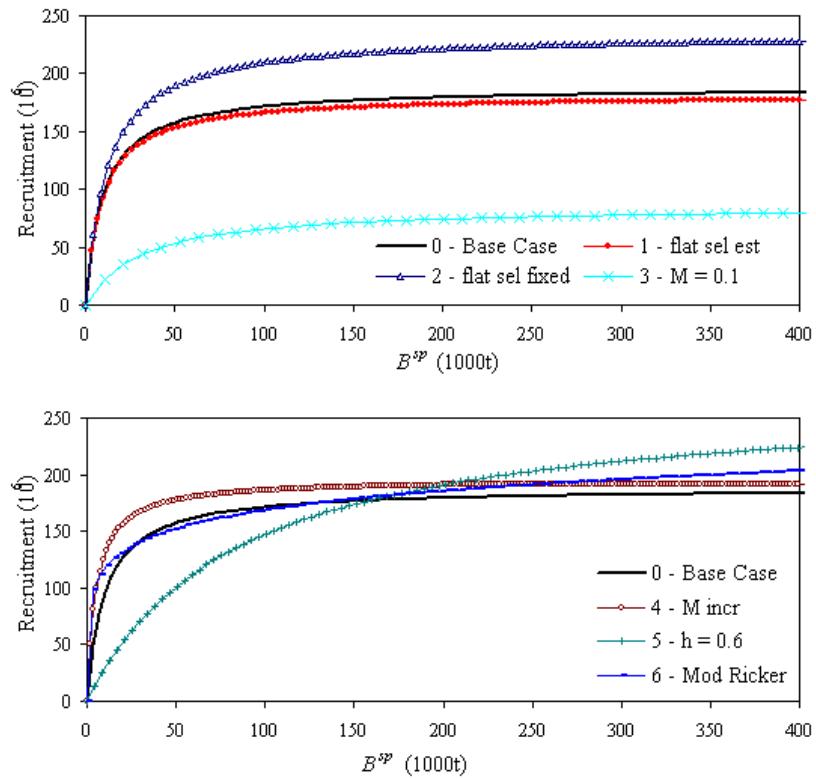


Fig. 2: Stock-recruitment relationships for a series of SCAA variants.

APPENDIX A – Data

Table A1: Landings (tons) for Greenland Halibut in Sub-area 2 and Div. 3KLMNO (Healey and Mahé, 2008).

Year	Landings (t)	Year	Landings (t)
1960	938	1984	26711
1961	741	1985	20347
1962	588	1986	17976
1963	1621	1987	32442
1964	4252	1988	19215
1965	10069	1989	20034
1966	19276	1990	47454
1967	26525	1991	65008
1968	32392	1992	63193
1969	37275	1993	62455
1970	36889	1994	51029
1971	24834	1995	15272
1972	30038	1996	18840
1973	29105	1997	19858
1974	27588	1998	19946
1975	28814	1999	24226
1976	24611	2000	34177
1977	32048	2001	38232
1978	39070	2002	34062
1979	34104	2003	35151
1980	32867	2004	25486
1981	30754	2005	23255
1982	26278	2006	23531
1983	27861	2007	22747

Table A2. Catch at age matrix (000s) for Greenland Halibut in Sub-Area 2 and Divisions 3KLMNO (Healey and Mahé, 2008).

Year	1	2	3	4	5	6	7	8	9	10	11	12	13	14+
1975	0	0	0	0	334	2819	5750	4956	3961	1688	702	135	279	288
1976	0	0	0	0	17	610	3231	5413	3769	2205	829	260	101	53
1977	0	0	0	0	534	5012	10798	7346	2933	1013	220	130	116	84
1978	0	0	0	0	2982	8415	8970	7576	2865	1438	723	367	222	258
1979	0	0	0	0	2386	8727	12824	6136	1169	481	287	149	143	284
1980	0	0	0	0	209	2086	9150	9679	5398	3828	1013	128	53	27
1981	0	0	0	0	863	4517	9806	11451	4307	890	256	142	43	69
1982	0	0	0	0	269	2299	6319	5763	3542	1684	596	256	163	191
1983	0	0	0	0	701	3557	9800	7514	2295	692	209	76	106	175
1984	0	0	0	0	902	2324	5844	7682	4087	1259	407	143	106	183
1985	0	0	0	0	1983	5309	5913	3500	1380	512	159	99	87	86
1986	0	0	0	0	280	2240	6411	5091	1469	471	244	140	70	117
1987	0	0	0	0	137	1902	11004	8935	2835	853	384	281	225	349
1988	0	0	0	0	296	3186	8136	4380	1288	465	201	105	107	129
1989	0	0	0	0	181	1988	7480	4273	1482	767	438	267	145	71
1990	0	0	0	95	1102	6758	12632	7557	4072	2692	1204	885	434	318
1991	0	0	0	220	2862	7756	13152	10796	7145	3721	1865	1216	558	422
1992	0	0	0	1064	4180	10922	20639	12205	4332	1762	1012	738	395	335
1993	0	0	0	1010	9570	15928	17716	11918	4642	1836	1055	964	401	182
1994	0	0	0	5395	16500	15815	11142	6739	3081	1103	811	422	320	215
1995	0	0	0	323	1352	2342	3201	2130	1183	540	345	273	251	201
1996	0	0	0	190	1659	5197	6387	1914	956	504	436	233	143	89
1997	0	0	0	335	1903	4169	7544	3215	1139	606	420	246	137	89
1998	0	0	0	552	3575	5407	5787	3653	1435	541	377	161	92	51
1999	0	0	0	297	2149	5625	8611	3793	1659	623	343	306	145	151
2000	0	0	0	271	2029	12583	21175	3299	973	528	368	203	129	104
2001	0	0	0	448	2239	12163	22122	5154	1010	495	439	203	156	75
2002	0	0	0	479	1662	7239	17581	6607	1244	659	360	224	126	81
2003	0	0	0	1279	4491	10723	16764	6385	1614	516	290	144	76	85
2004	0	0	0	897	4062	8236	10542	4126	1307	529	289	184	87	75
2005	0	0	0	534	1652	5999	10313	3996	1410	444	244	114	64	46
2006	0	0	0	216	1869	6450	12144	4902	1089	372	136	47	32	40
2007	0	0	0	88	570	3732	11912	5414	1230	472	163	80	41	29

Table A3. Catch weights-at-age (kg) matrix for Greenland Halibut in Sub-Area 2 and Divisions 3KLMNO (Healy and Mahé, 2008).

Year	1	2	3	4	5	6	7	8	9	10	11	12	13	14+
1975	0.000	0.000	0.126	0.244	0.609	0.760	0.955	1.190	1.580	2.210	2.700	3.370	3.880	5.764
1976	0.000	0.000	0.126	0.244	0.609	0.760	0.955	1.190	1.580	2.210	2.700	3.370	3.880	5.144
1977	0.000	0.000	0.126	0.244	0.609	0.760	0.955	1.190	1.580	2.210	2.700	3.370	3.880	5.992
1978	0.000	0.000	0.126	0.244	0.609	0.760	0.955	1.190	1.580	2.210	2.700	3.370	3.880	5.894
1979	0.000	0.000	0.126	0.244	0.609	0.760	0.955	1.190	1.580	2.210	2.700	3.370	3.880	6.077
1980	0.000	0.000	0.126	0.244	0.514	0.659	0.869	1.050	1.150	1.260	1.570	2.710	3.120	5.053
1981	0.000	0.000	0.126	0.244	0.392	0.598	0.789	0.985	1.240	1.700	2.460	3.510	4.790	7.426
1982	0.000	0.000	0.126	0.244	0.525	0.684	0.891	1.130	1.400	1.790	2.380	3.470	4.510	7.359
1983	0.000	0.000	0.126	0.244	0.412	0.629	0.861	1.180	1.650	2.230	3.010	3.960	5.060	7.061
1984	0.000	0.000	0.126	0.244	0.377	0.583	0.826	1.100	1.460	1.940	2.630	3.490	4.490	7.016
1985	0.000	0.000	0.126	0.244	0.568	0.749	0.941	1.240	1.690	2.240	2.950	3.710	4.850	7.010
1986	0.000	0.000	0.126	0.244	0.350	0.584	0.811	1.100	1.580	2.120	2.890	3.890	4.950	7.345
1987	0.000	0.000	0.126	0.244	0.364	0.589	0.836	1.160	1.590	2.130	2.820	3.600	4.630	6.454
1988	0.000	0.000	0.126	0.244	0.363	0.569	0.805	1.163	1.661	2.216	3.007	3.925	5.091	7.164
1989	0.000	0.000	0.126	0.244	0.400	0.561	0.767	1.082	1.657	2.237	2.997	3.862	4.919	6.370
1990	0.000	0.000	0.090	0.181	0.338	0.546	0.766	1.119	1.608	2.173	2.854	3.731	4.691	6.391
1991	0.000	0.000	0.126	0.244	0.383	0.592	0.831	1.228	1.811	2.461	3.309	4.142	5.333	7.081
1992	0.000	0.000	0.175	0.289	0.430	0.577	0.793	1.234	1.816	2.462	3.122	3.972	5.099	6.648
1993	0.000	0.000	0.134	0.232	0.368	0.547	0.809	1.207	1.728	2.309	2.999	3.965	4.816	6.489
1994	0.000	0.000	0.080	0.196	0.330	0.514	0.788	1.179	1.701	2.268	2.990	3.766	4.882	6.348
1995	0.000	0.000	0.080	0.288	0.363	0.531	0.808	1.202	1.759	2.446	3.122	3.813	4.893	6.790
1996	0.000	0.000	0.161	0.242	0.360	0.541	0.832	1.272	1.801	2.478	3.148	3.856	4.953	6.312
1997	0.000	0.000	0.120	0.206	0.336	0.489	0.771	1.159	1.727	2.355	3.053	3.953	5.108	6.317
1998	0.000	0.000	0.119	0.228	0.373	0.543	0.810	1.203	1.754	2.351	3.095	4.010	5.132	6.124
1999	0.000	0.000	0.176	0.253	0.358	0.533	0.825	1.253	1.675	2.287	2.888	3.509	4.456	5.789
2000	0.000	0.000	0.000	0.254	0.346	0.524	0.787	1.192	1.774	2.279	2.895	3.645	4.486	5.531
2001	0.000	0.000	0.000	0.249	0.376	0.570	0.830	1.168	1.794	2.367	2.950	3.715	4.585	5.458
2002	0.000	0.000	0.217	0.251	0.369	0.557	0.841	1.193	1.760	2.277	2.896	3.579	4.407	5.477
2003	0.000	0.000	0.188	0.247	0.389	0.564	0.822	1.199	1.651	2.166	2.700	3.404	4.377	5.409
2004	0.000	0.000	0.180	0.249	0.376	0.535	0.808	1.196	1.629	2.146	2.732	3.538	4.381	5.698
2005	0.000	0.000	0.252	0.301	0.396	0.564	0.849	1.247	1.691	2.177	2.705	3.464	4.264	5.224
2006	0.000	0.000	0.129	0.267	0.405	0.605	0.815	1.092	1.495	1.874	2.396	3.139	3.747	4.701
2007	0.000	0.000	0.000	0.276	0.389	0.581	0.833	1.137	1.500	1.948	2.607	3.057	3.869	4.954

Table A4: Proportion mature-at-age for Greenland Halibut in Sub-Area 2 and Divisions 3KLMNO (Healy pers. comm.). Note in the assessment, the maturity-at-age in 2008 and pre-1975 is taken as the average over the 1975-2007 period.

	0	1	2	3	4	5	6	7	8	9	10	11	12	13	14+
1975	0.00	0.00	0.00	0.00	0.01	0.00	0.01	0.04	0.04	0.03	0.12	0.21	0.34	0.50	0.77
1976	0.00	0.00	0.00	0.00	0.00	0.01	0.01	0.02	0.06	0.07	0.06	0.21	0.34	0.50	0.72
1977	0.00	0.00	0.00	0.00	0.00	0.00	0.02	0.01	0.04	0.11	0.12	0.14	0.34	0.50	0.79
1978	0.00	0.00	0.00	0.00	0.00	0.01	0.01	0.04	0.03	0.08	0.18	0.20	0.29	0.50	0.78
1979	0.00	0.00	0.00	0.00	0.00	0.01	0.01	0.02	0.06	0.06	0.16	0.28	0.31	0.50	0.80
1980	0.00	0.00	0.00	0.00	0.00	0.01	0.02	0.04	0.11	0.12	0.28	0.41	0.45	0.76	
1981	0.00	0.00	0.00	0.00	0.00	0.01	0.02	0.03	0.07	0.18	0.23	0.45	0.55	0.76	
1982	0.00	0.00	0.00	0.00	0.00	0.00	0.01	0.04	0.06	0.13	0.28	0.40	0.63	0.77	
1983	0.00	0.00	0.00	0.00	0.00	0.00	0.01	0.02	0.07	0.12	0.24	0.40	0.59	0.80	
1984	0.00	0.00	0.00	0.00	0.00	0.00	0.00	0.02	0.04	0.11	0.21	0.38	0.54	0.84	
1985	0.00	0.00	0.00	0.00	0.00	0.00	0.00	0.01	0.03	0.07	0.19	0.35	0.56	0.78	
1986	0.00	0.00	0.00	0.00	0.00	0.00	0.00	0.00	0.01	0.02	0.05	0.13	0.30	0.51	0.79
1987	0.00	0.00	0.00	0.00	0.00	0.00	0.00	0.00	0.02	0.06	0.10	0.22	0.43	0.77	
1988	0.00	0.00	0.00	0.00	0.00	0.00	0.00	0.01	0.00	0.01	0.04	0.15	0.17	0.34	0.71
1989	0.00	0.00	0.00	0.00	0.00	0.00	0.00	0.02	0.00	0.02	0.09	0.33	0.29	0.57	
1990	0.00	0.00	0.00	0.00	0.00	0.00	0.00	0.01	0.05	0.16	0.08	0.21	0.58	0.52	
1991	0.00	0.00	0.00	0.00	0.00	0.00	0.00	0.01	0.02	0.15	0.97	0.25	0.41	0.74	
1992	0.00	0.00	0.00	0.00	0.00	0.00	0.00	0.01	0.03	0.05	0.38	1.00	0.56	0.73	
1993	0.00	0.00	0.00	0.00	0.00	0.00	0.00	0.00	0.02	0.10	0.11	0.68	1.00	0.84	
1994	0.00	0.00	0.00	0.00	0.00	0.00	0.00	0.01	0.01	0.06	0.26	0.25	0.88	0.99	
1995	0.00	0.00	0.00	0.00	0.00	0.00	0.00	0.01	0.02	0.03	0.17	0.53	0.47	0.98	
1996	0.00	0.00	0.00	0.00	0.00	0.00	0.00	0.01	0.02	0.04	0.08	0.36	0.78	0.80	
1997	0.00	0.00	0.00	0.00	0.00	0.00	0.00	0.00	0.01	0.04	0.10	0.20	0.61	0.91	
1998	0.00	0.00	0.00	0.00	0.00	0.00	0.00	0.00	0.00	0.04	0.10	0.21	0.43	0.86	
1999	0.00	0.00	0.00	0.00	0.00	0.00	0.00	0.00	0.01	0.02	0.09	0.21	0.41	0.80	
2000	0.00	0.00	0.00	0.00	0.00	0.00	0.00	0.01	0.01	0.04	0.12	0.21	0.41	0.73	
2001	0.00	0.00	0.00	0.00	0.00	0.00	0.00	0.00	0.03	0.07	0.18	0.53	0.41	0.69	
2002	0.00	0.00	0.00	0.00	0.00	0.00	0.00	0.00	0.02	0.07	0.36	0.53	0.90	0.71	
2003	0.00	0.00	0.00	0.00	0.00	0.00	0.00	0.00	0.02	0.06	0.17	0.82	0.85	0.93	
2004	0.00	0.00	0.00	0.00	0.00	0.00	0.00	0.00	0.02	0.06	0.24	0.35	0.97	0.97	
2005	0.00	0.00	0.00	0.00	0.00	0.00	0.00	0.00	0.02	0.06	0.24	0.56	0.58	1.00	
2006	0.00	0.00	0.00	0.00	0.00	0.00	0.00	0.00	0.02	0.06	0.24	0.56	0.80	0.86	
2007	0.00	0.00	0.00	0.00	0.00	0.00	0.00	0.00	0.02	0.06	0.24	0.49	0.79	0.94	

Table A6: Survey data in terms of weight for ages combined: 2J3K Fall and 3LNO Spr (Healey, 2008), EU survey (Vázquez and González-Troncoso, 2008).

Year	2J3K Fall	EU survey	3LNO - Spr
	Mean weight (kg)/tow	Biomass (tons)	Mean weight (kg)/tow
1978	38.4		
1979	28.1		
1980	30.0		
1981	32.1		
1982	35.6		
1983	36.9		
1984	37.2		
1985	27.5		
1986	35.4		
1987	25.5		
1988	23.6	6926	
1989	25.4	4472	
1990	21.2	5799	
1991	11.5	8169	
1992	8.2	8728	
1993	15.3	6529	
1994	10.8	8037	
1995	14.1	10875	
1996	21.6	11594	1.43
1997	24.8	16098	2.10
1998	23.8	24229	3.50
1999	32.5	21207	2.33
2000	23.9	16959	2.30
2001	22.7	13872	1.13
2002	14.1	12100	0.53
2003	15.3	6214	1.13
2004	17.5	12292	0.87
2005	20.3	11698	1.23
2006	25.7	11706	
2007	29.1	13040	2.17

Appendix B - The SCAA Model

The model used for these assessments is an Age-Structured Production Model (ASPM). Models of this type fall within the more general class of Statistical Catch-at-Age Analyses. The approach used in an ASPM assessment involves constructing an age-structured model of the population dynamics and fitting it to the available abundance indices by maximising the likelihood function. The model equations and the general specifications of the model are described below, followed by details of the contributions to the (penalised) log-likelihood function from the different sources of data available and assumptions concerning the stock-recruitment relationship. Quasi-Newton minimization is used to minimize the total negative log-likelihood function (the package AD Model Builder™, Otter Research, Ltd is used for this purpose).

B.1. Population dynamics

B.1.1 Numbers-at-age

The resource dynamics are modelled by the following set of population dynamics equations:

$$N_{y+1,1} = R_{y+1} \quad (\text{B1})$$

$$N_{y+1,a+1} = (N_{y,a} e^{-M_a/2} - C_{y,a}) e^{-M_a/2} \quad \text{for } 1 \leq a \leq m-2 \quad (\text{B2})$$

$$N_{y+1,m} = (N_{y,m-1} e^{-M_{m-1}/2} - C_{y,m-1}) e^{-M_{m-1}/2} + (N_{y,m} e^{-M_m/2} - C_{y,m}) e^{-M_m/2} \quad (\text{B3})$$

where

$N_{y,a}$ is the number of fish of age a at the start of year y (which refers to a calendar year),

R_y is the recruitment (number of 0-year-old fish) at the start of year y ,

M_a denotes the natural mortality rate for fish of age a ,

$C_{y,a}$ is the predicted number of fish of age a caught in year y , and

m is the maximum age considered (taken to be a plus-group).

These equations reflect Pope's form of the catch equation (Pope, 1972) (the catches are assumed to be taken as a pulse in the middle of the year) rather than the more customary Baranov form (Baranov, 1918) (for which catches are incorporated under the assumption of steady continuous fishing mortality). Pope's form has been used in order to simplify computations. As long as mortality rates are not too high, the differences between the Baranov and Pope formulations will be minimal.

B.1.2. Recruitment

The number of recruits at the start of year y is assumed to be related to the spawning stock size (i.e. the biomass of mature fish) by a Beverton-Holt stock-recruitment relationship (Beverton and Holt, 1957), parameterised in terms of the "steepness" of the stock-recruitment relationship, h , and the pre-exploitation equilibrium spawning biomass, K^{sp} , and recruitment, R_0 and allowing for annual fluctuation about the deterministic relationship:

$$R_y = \frac{4hR_0B_y^{sp}}{K^{sp}(1-h)+(5h-1)B_y^{sp}} e^{(\zeta_y - \sigma_K^2/2)} \quad (\text{B4})$$

where

ξ_y reflects fluctuation about the expected recruitment for year y , which is assumed to be normally distributed with standard deviation σ_R (which is input in the applications considered here); these residuals are treated as estimable parameters in the model fitting process.

B_y^{sp} is the spawning biomass at the start of year y , computed as:

$$B_y^{sp} = \sum_{a=1}^m f_{y,a} w_{y,a}^{strt} N_{y,a} \quad (\text{B5})$$

where

$w_{y,a}^{strt}$ is the mass of fish of age a during spawning, and

$f_{y,a}$ is the proportion of fish of age a that are mature.

In the fitting procedure, K^{sp} is estimated while h has thus far been fixed at 0.9 for reasons elaborated in the main text.

B.1.3. Total catch and catches-at-age

The catch by mass in year y is given by:

$$C_y = \sum_{a=1}^m w_{y,a}^{mid} C_{y,a} = \sum_{a=1}^m w_{y,a}^{mid} N_{y,a} e^{-M_a/2} S_{y,a} F_y \quad (\text{B6})$$

where

$w_{y,a}^{mid}$ denotes the mass of fish of age a landed in year y ,

$C_{y,a}$ is the catch-at-age, i.e. the number of fish of age a , caught in year y ,

$S_{y,a}$ is the commercial selectivity (i.e. combination of availability and vulnerability to fishing gear) at age a for year y ; when $S_{y,a} = 1$, the age-class a is said to be fully selected, and

F_y is the proportion of a fully selected age class that is fished.

The model estimate of the mid-year exploitable (“available”) component of biomass is calculated by converting the numbers-at-age into mid-year mass-at-age (using the individual weights of the landed fish) and applying natural and fishing mortality for half the year:

$$B_y^{ex} = \sum_{a=1}^m w_{y,a}^{mid} S_{y,a} N_{y,a} e^{-M_a/2} (1 - S_{y,a} F_y / 2) \quad (\text{B7})$$

whereas for survey estimates of biomass in spring:

$$B_y^{surv,spring} = \sum_{a=1}^m w_{y,a}^{mid} S_a^{surv} N_{y,a} e^{-M_a/4} (1 - S_{y,a} F_y / 4) \quad (\text{B8})$$

Summer:

$$B_y^{surv,summer} = \sum_{a=1}^m w_{y,a}^{mid} S_a^{surv} N_{y,a} e^{-M_a/2} (1 - S_{y,a} F_y / 2) \quad (\text{B9})$$

and fall:

$$B_y^{surv,fall} = \sum_{a=1}^m w_{y,a}^{mid} S_a^{surv} N_{y,a} e^{-M_a 3/4} (1 - S_{y,a} F_y 3/4) \quad (\text{B10})$$

where

S_a^{surv} is the survey selectivity for age a (which is sometimes generalised to be year-dependent).

B.1.4. Initial conditions

For the first year (y_0) considered in the model therefore, the stock is assumed to be at a fraction (θ) of its pre-exploitation biomass, i.e.:

$$B_{y_0}^{sp} = \theta \cdot K^{sp} \quad (\text{B11})$$

with the starting age structure:

$$N_{y_0,a} = R_{start} N_{start,a} \quad \text{for } 1 \leq a \leq m \quad (\text{B12})$$

where

$$N_{start,1} = 1 \quad (\text{B13})$$

$$N_{start,a} = N_{start,a-1} e^{-M_{a-1}} (1 - \phi S_{a-1}) \quad \text{for } 2 \leq a \leq m-1 \quad (\text{B14})$$

$$N_{start,m} = N_{start,m-1} e^{-M_{m-1}} (1 - \phi S_{m-1}) / (1 - e^{-M_m} (1 - \phi S_m)) \quad (\text{B15})$$

where ϕ characterises the average fishing proportion over the years immediately preceding y_0 .

Unless indicated otherwise though, the stock is assumed to be at pristine equilibrium in 1960, i.e. $\theta=1$ and $\phi=0$ for the results reported here.

B.2. The (penalised) likelihood function

The model can be fit to (a subset of) CPUE and survey abundance indices, and commercial and survey catch-at-age data to estimate model parameters (which may include residuals about the stock-recruitment function, the fishing selectivities, the annual catches or natural mortality, facilitated through the incorporation of penalty functions described below). Contributions by each of these to the negative of the (penalised) log-likelihood (-lnL) are as follows.

B.2.1 CPUE relative abundance data

The likelihood is calculated assuming that an observed CPUE index for a particular fishing fleet is log-normally distributed about its expected value:

$$I_y^i = \hat{I}_y^i \exp(\varepsilon_y^i) \quad \text{or} \quad \varepsilon_y^i = \ell \ln(I_y^i) - \ell \ln(\hat{I}_y^i) \quad (\text{B16})$$

where

I_y^i is the CPUE index for year y and series i ,

$\hat{I}_y^i = \hat{q}^i \hat{B}_y^{ex}$ is the corresponding model estimate, where \hat{B}_y^{ex} is the model estimate of exploitable resource biomass, given by equation (B7)¹,

\hat{q}^i is the constant of proportionality (catchability) for CPUE series i , and

$$\varepsilon_y^i \text{ from } N\left(0, (\sigma_y^i)^2\right).$$

The contribution of the CPUE data to the negative of the log-likelihood function (after removal of constants) is then given by:

$$-\ln L^{CPUE} = \sum_i \sum_y \left[\ln(\sigma_y^i) + (\varepsilon_y^i)^2 / 2(\sigma_y^i)^2 \right] \quad (B17)$$

where

σ_y^i is the standard deviation of the residuals for the logarithm of index i in year y .

Homoscedasticity of residuals is assumed, so that $\sigma_y^i = \sigma^i$ is estimated in the fitting procedure by its maximum likelihood value:

$$\hat{\sigma}^i = \sqrt{1/n_i \sum_y (\ln(I_y^i) - \ln(q^i \hat{B}_y^{ex}))^2} \quad (B18)$$

where

n_i is the number of data points for CPUE index i .

The catchability coefficient q^i for CPUE index i is estimated by its maximum likelihood value:

$$\ln \hat{q}^i = 1/n_i \sum_y (\ln I_y^i - \ln \hat{B}_y^{ex}) \quad (B19)$$

B.2.2. Survey abundance data

In general, data from the surveys are treated as relative abundance indices in exactly the same manner to the CPUE series above, with survey selectivity function S_a^{surv} replacing the commercial selectivity $S_{y,a}$. Account is also taken of the time of year when the survey is held. For these analyses, selectivities are estimated as detailed in section B.4.2 below.

To allow for serial correlation between the survey residuals, the ε_y^i input to equation B17 is given by:

$$\varepsilon_y^i = \lambda_y^i - \rho \lambda_{y-1}^i \quad (B20)$$

where

$$\lambda_y^i = \ln(I_y^i) - \ln(\hat{I}_y^i)$$

¹ Ideally \hat{B}_y^{ex} should be fleet specific, corresponding to the selectivity for the fleet linked to CPUE index i . However, this requires the total annual catch and catch-at-age data to be provided on a fleet-disaggregated basis, and these data are not immediately available in this form.

ρ is the serial correlation coefficient, which is estimated (or set to zero in the case of the Baseline assessment B1). Note that ρ could be series dependent, but analyses for the data set available indicated that estimation of series-specific values was not justified in AIC terms. The standard deviation of the ε_y^i is termed σ_{surv} in Table 1.

B.2.3. Commercial catches-at-age

The contribution of the catch-at-age data to the negative of the log-likelihood function under the assumption of an “adjusted” lognormal error distribution is given by:

$$-\ln L^{CAA} = \sum_y \sum_a \left[\ln \left(\sigma_{\text{com}} / \sqrt{p_{y,a}} \right) + p_{y,a} (\ln p_{y,a} - \ln \hat{p}_{y,a})^2 / 2(\sigma_{\text{com}})^2 \right] \quad (\text{B21})$$

where

$p_{y,a} = C_{y,a} / \sum_{a'} C_{y,a'}$ is the observed proportion of fish caught in year y that are of age a ,

$\hat{p}_{y,a} = \hat{C}_{y,a} / \sum_{a'} \hat{C}_{y,a'}$ is the model-predicted proportion of fish caught in year y that are of age a ,

where

$$\hat{C}_{y,a} = N_{y,a} e^{-M_a/2} S_{y,a} F_y \quad (\text{B22})$$

and

σ_{com} is the standard deviation associated with the catch-at-age data (termed “ σ_{comCAA} ” in Table 1), which is estimated in the fitting procedure by:

$$\hat{\sigma}_{\text{com}} = \sqrt{\sum_y \sum_a p_{y,a} (\ln p_{y,a} - \ln \hat{p}_{y,a})^2 / \sum_y \sum_a 1} \quad (\text{B23})$$

The log-normal error distribution underlying equation (B21) is chosen on the grounds that (assuming no ageing error) variability is likely dominated by a combination of interannual variation in the distribution of fishing effort, and fluctuations (partly as a consequence of such variations) in selectivity-at-age, which suggests that the assumption of a constant coefficient of variation is appropriate. However, for ages poorly represented in the sample, sampling variability considerations must at some stage start to dominate the variance. To take this into account in a simple manner, motivated by binomial distribution properties, the observed proportions are used for weighting so that undue importance is not attached to data based upon a few samples only.

Commercial catches-at-age are incorporated in the likelihood function using equation (B21), for which the summation over age a is taken from age a_{minus} (considered as a minus group) to a_{plus} (a plus group).

B.2.4. Survey catches-at-age

The survey catches-at-age are incorporated into the negative of the log-likelihood in an analogous manner to the commercial catches-at-age, assuming an adjusted log-normal error distribution (equation (B21)) where:

$p_{y,a} = C_{y,a}^{\text{surv}} / \sum_{a'} C_{y,a'}^{\text{surv}}$ is the observed proportion of fish of age a in year y ,

$\hat{p}_{y,a}$ is the expected proportion of fish of age a in year y in the survey, given by:

$$\hat{p}_{y,a} = S_a^{\text{surv}} N_{y,a} / \sum_{a'=0}^m S_{a'}^{\text{surv}} N_{y,a'} \quad \text{for begin-year surveys.} \quad (\text{B24})$$

The residual standard deviation (analogous to σ_{com} , alternatively termed “ σ_{comCAA} ” in the previous section, is termed $\sigma_{survCAA}$ in Table 1.

B.2.5. Stock-recruitment function residuals

The stock-recruitment residuals are assumed to be log-normally distributed. Thus, the contribution of the recruitment residuals to the negative of the (now penalised) log-likelihood function is given by:

$$-\ell nL^{SRpen} = \sum_{y=y1}^{y2} [\varepsilon_y^2 / 2\sigma_R^2] \quad (\text{B25})$$

where

ε_y from $N(0, (\sigma_R)^2)$, which is estimated for year $y1$ to $y2$ (see equation (B4)), and

σ_R is the standard deviation of the log-residuals, which is input.

B.2.6. Selectivity residuals

In some instances, variations around the fishing selectivity functions are estimated in two-year periods:

$$S_a \rightarrow S_{y,a} = S_a e^{\Omega_{y,a}} \quad (\text{B26})$$

The contribution of the selectivity residuals to the negative of the penalised log-likelihood is given by:

$$-\ell nL^{Selpen} = \sum_{y=y1}^{y2} \sum_{a=a1}^{a2} [\Omega_{y,a}^2 / 2\sigma_\Omega^2] \quad (\text{B27})$$

where

$\Omega_{y,a}$ from $N(0, (\sigma_\Omega)^2)$, which is estimated for year $y1$ to $y2$ and age $a1$ to $a2$, and

σ_Ω is the standard deviation of the residuals, which is input.

B.2.7. Annual catch residuals

In some instances, differences between the reported and the actual annual catches are estimated.

$$C_y \rightarrow C_y^{reported} = C_y^{actual} e^{\omega_y} \quad (\text{B28})$$

The contribution of the catch residuals to the negative of the penalised log-likelihood is given by:

$$-\ell nL^{Cpen} = \sum_{y=y1}^{y2} [\omega_y^2 / 2\sigma_C^2] \quad (\text{B29})$$

where

ω_y from $N(0, (\sigma_C)^2)$, which is estimated for year $y1$ to $y2$, and

σ_C is the standard deviation of the residuals, which is input.

B.2.7. Mortality residuals

In some instances, variations about the default age- and year-independent natural mortality are estimated.

$$M_a \rightarrow M_a e^{\zeta_{y,a}} \quad (\text{B30})$$

The contribution of the catch residuals to the negative of the penalised log-likelihood is given by:

$$-\ell n L^{M\text{pen}} = \sum_{y=y1}^{y2} \sum_{a=a1}^{a2} [\zeta_{y,a}^2 / 2\sigma_M^2] \quad (\text{B31})$$

where

$\zeta_{y,a}$ from $N(0, (\sigma_M)^2)$, which is estimated for year $y1$ to $y2$, and age $a1$ to $a2$, and

σ_M is the standard deviation of the residuals, which is input.

B.3. Estimation of precision

Where quoted, CVs are Hessian-based.

B.4. Model parameters

B.4.1. Fishing selectivity-at-age:

The commercial fishing selectivity, S_a , is estimated separately for ages 5-11. The estimated decreases from ages 6 to 5 and ages 11 to 12 are assumed to continue exponentially to ages 0 and 14+ respectively. Similarly, the selectivities for the surveys are estimate separately for ages 1-11 for the Canadian Fall and EU surveys and for ages 1-8 for the Canadian spring surveys. The estimated decreases from ages 2 to 1 and from ages 10 to 11 (7 to 8 for the Canadian spring survey) are assumed to continue exponentially to ages 0 and 14+ respectively.

B.4.2. Other parameters

Plus group:	
m	14
Commercial CAA:	
a_{minus}	5
a_{plus}	12
Survey CAA:	
a_{minus}	1
a_{plus}	11/8
Stock-recruitment residuals:	
Δ_R	0.25
y_1	1961
y_2	2007
Natural mortality:	
M	age independent, fixed
Age-at-maturity:	
$f_{y,a}$	input, see Table A4
Weight-at-age:	
$w_{y,a}$	input, same for start-and mid-year, see Table A3
Initial conditions:	
Δ	1 (unless otherwise specified)
Δ	0 (unless otherwise specified)
Survey serial correlation:	
Δ	0 (unless otherwise specified)

Addendum

In the calculations reported here, the following modifications to the specifications above were incorporated:

- 1) Taking the modelled population age structure to a plus-group age m of 20 rather than 14 so that any decreasing selectivity trend continues to a larger age. For the commercial selectivity, the estimated decrease from ages 11 to 12 are assumed to continue exponentially to age 20+ for the Reference Case. Similarly for the survey selectivities, the estimated decrease from ages 10 to 11 (7 to 8 for the Canadian spring survey) is assumed to continue exponentially to age 20+.
- 2) The inclusion of correlation, in both year and age, in survey proportions-at-age data. To allow for serial correlation between the survey proportion-at-age residuals, equation B21 above is replaced by:

$$-\ln L^{CAA} = \sum_{i=1}^{n_{surv}} \left\{ \begin{aligned} & \sum_{y=2}^n \sum_{a=a_{min_us}+1}^{a_{plus}} \left[\ln \left(\sigma_{surv}^i / \sqrt{p_{y,a}^i} \right) + p_{y,a}^i (\eta_{y,a}^i)^2 / 2(\sigma_{surv}^i)^2 \right] \\ & + \sum_{a=a_{min_us}}^{a_{plus}} \left(\sum_{y=1}^n p_{y,a}^i - \sum_{y=1}^n \bar{p}_{y,a}^i \right)^2 \end{aligned} \right\} \quad (B32)$$

with

$$\eta_{y,a}^i = (\mathcal{E}_{y,a}^i - \rho_{CAAage}^i \mathcal{E}_{y,a-1}^i) - \rho_{CAAYr}^i (\mathcal{E}_{y-1,a}^i - \rho_{CAAage}^i \mathcal{E}_{y-1,a-1}^i) \quad (B33)$$

$$\mathcal{E}_{y,a}^i = \sqrt{p_{y,a}^i} (\ln p_{y,a}^i - \ln \hat{p}_{y,a}^i) \quad (B34)$$

and

$$\hat{\sigma}_{surv}^i = \sqrt{\sum_y \sum_a \eta_{y,a}^i / \sum_y \sum_a 1} \quad (B35)$$

ρ_{CAAage}^i is the age serial correlation coefficient for survey series i , which is estimated, and

ρ_{CAAYr}^i is the year serial correlation coefficient for survey series i , which is estimated.

The second term in the likelihood has been added so that the average predicted proportion-at-age is close to that observed. This is necessary here because taking account of serial correlation loses the information otherwise in the likelihood to fit to data for the youngest age-class.